

Competition Policy and Productivity Growth: An Empirical Assessment*

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Abstract

This paper empirically investigates the effectiveness of competition policy by estimating its impact on Total Factor Productivity (TFP) growth for 22 industries in 12 OECD countries over the period 1995-2005. We find a robust positive and significant effect of competition policy as measured by newly created indexes. We provide several arguments and results based on instrumental variables estimators as well as non-linearities to support the claim that the established link can be interpreted in a causal way. At a disaggregated level, the effect on TFP growth is particularly strong for specific aspects of competition policy related to its institutional set up and antitrust activities (rather than merger control). The effect is strengthened by good legal systems, suggesting complementarities between competition policy and the efficiency of law enforcement institutions.

Keywords: Competition Policy, Productivity Growth, TFP, Institutions, Deterrence, OECD
JEL classification: L4, K21, O4, C23

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1 Introduction

The aim of this paper is to assess the effectiveness of competition policy in providing higher welfare to society thanks to improved efficiency and productivity.¹ While most economists, starting from Adam Smith, agree that *competition* works in the general interest, there is no such consensus on the ability of *competition policy* to be socially beneficial. Some economists, dating back to the "Austrian School" (e.g. Von Mises, 1940), argue that any state intervention that interferes with free markets will make society worse off. According to them, competition policy is not an exception, even though its aim is to safeguard effective competition.

More recently, Crandall and Winston (2003) claimed that, at least in the US, antitrust law has been ineffective. They maintain that its poor performance is mostly due to the difficulty of distinguishing genuine and healthy competition from anti-competitive behaviors (in all areas of competition law) and to the undervalued power of the markets to curb anti-competitive abuses. They do not ask for a repeal of antitrust law, but urge applying it only for blatant price-fixing and merger to monopoly. Baker (2003) and Werden (2003) disagreed with Crandall and Winston's point of view. They argue that the net effect of competition policy on social welfare is positive. In their opinion, competition policy improves social welfare by also (or mostly) inducing firms to forgo anti-competitive behaviors without an explicit intervention of any competition authority, i.e. by deterring them. The debate appears to be still unsettled. As noted by Whinston (2006), even in the most established area of competition policy, cartel deterrence, 'strong' empirical evidence of the actual effects of the practices forbidden by antitrust law (e.g. competitors communicating on prices), and of active antitrust law enforcement on social welfare, is still missing.

This paper is an attempt to provide 'strong' empirical evidence, at least with respect to the effectiveness of the application of competition law in general. In order to do so, we estimate the impact of competition policy and some of its components on total factor productivity (TFP) growth on a sample of 22 industries in 12 OECD countries over the period 1995-2005. To mea-

¹By competition policy we mean the set of prohibitions and obligations that forms the substantive rules of competition (or antitrust) law together with the array of tools available to competition authorities for policing and punishing any violation of the same rules.

sure competition policy, we identify a set of its institutional and enforcement features that we consider to be key in deterring anti-competitive behavior. We then aggregate these variables to form a set of summary indicators, the Competition Policy Indicators (CPIs). We generate an Aggregate CPI that summarizes all the key features of the competition policy of a country, as well as more disaggregated ones that refer only to the features of competition policy relative to specific behaviors (i.e. cartels, other competitive agreements and abuses of dominance – collectively referred to as ‘antitrust’ – and mergers), or only to the ‘institutional’ or the ‘enforcement’ features of a competition policy.

As a measure of efficiency we use TFP growth, which measures the increase in the amount of output that can be produced with a given increase in the amount of inputs. The theoretical and empirical literature has shown the existence of a positive relationship between competition and productivity. For this and related reasons, as we will argue in more detail in section 2, we believe that there should be a positive link between good competition policy and productivity. Since there is no consensus on the proper way to measure the extent of product-market competition and even the most widespread measure used to this end, the price-cost margin, has been strongly criticized (e.g. Boone, 2008), we think that to study the direct impact of competition policy on productivity growth, a crucial determinant of economic growth, is a proper way to measure the *gross* contribution of competition policy to social welfare.

In all specifications of our model, we control for country-industry and time fixed-effects, product market regulation, trade liberalization, and other likely determinants of productivity growth, and we find that the Aggregate CPI has a positive and highly significant effect on TFP growth. This impact is larger for industries far away from the technological frontier, suggesting that effective competition in such laggard sectors is even more important to foster productivity and increase efficiency. When we use the more disaggregated CPIs, separating the effects of the institutional and enforcement features, and distinguishing between mergers and antitrust, we find positive and significant coefficients estimates for all these indicators, though institutions and antitrust appear to have the strongest and a more significant impact on productivity growth. For the Aggregate CPI we find the same result both when we estimate the model by OLS, as well as in alternative IV specifications, which use either some political variables or

the competition policy in other countries as instruments for the policy. In addition to the IV estimation, we exploit the possible non-linearities in the effectiveness of competition policy on TFP growth to improve our identification strategy. Competition policy is expected to be more effective in countries with better legal institutions as well as in industries where no other sector-specific authorities are in charge of regulating the competitive processes. This is what we find.

The interaction between competition policy and institutions is not only part of our identification strategy. Indeed, competition policy does not work in isolation. Our CPIs describe some *internal* features of competition policy. However, the effectiveness of competition policy is also likely to depend on *external* factors: the quality of a country's institutions in general, and of its judicial system, in particular. These external factors may matter for two main reasons. First, the general quality of the institutions of a country creates an environment that affects the effectiveness of all public policies. In a context where public bodies in general are effective and efficient the bodies that preside over the enforcement of competition law also tend to be effective and efficient. Hence, if we do not control for institutions, the CPIs might capture some features that, instead, are a reflection of these more general factors. Second, inherent complementarities between competition policy and the judicial system might exist, as the enforcement of the policy is often done by the courts, directly or in appeal. For these reasons the courts, and the legal system in general, may play an important role in determining the deterrence properties of a competition policy regime. When we add the dimension of the quality of the institutions to our estimate, we observe that there are both direct effects of institutions on TFP growth and complementarities between them and our measures of competition policy. Indeed, we find that the effects of competition policy are strengthened in countries where the cost of enforcing contracts are low and the quality of the legal system is high, which points to sizable institutional complementarities between competition policy and the efficiency of legal institutions. These results suggest that competition policy grossly contributes to social welfare, especially in those countries where it is coupled with efficient and effective institutions.

The remainder of the paper is organized as follows. In Section 2 we briefly provide the theoretical background of our empirical research and relate our paper to the relevant literature.

Section 3 presents and discusses our empirical model and the identification strategy. Section 4 presents the data we use, the CPIs and how they have been built, and the political variables we use as instruments in the policy equation. Section 5 discusses our results and performs some robustness checks. Section 6 briefly concludes.

2 Theoretical Background and Literature Review

The objective of competition policy is to deter behaviors that reduce competition. Therefore, the causal link between competition policy and efficiency goes through the impact of the former on market competition. Aghion and Schankerman (2004) provide a theoretical framework for explaining this link. They point out that competition-enhancing policies may improve productivity by facilitating the weeding out of less efficient firms;² by promoting cost reduction investments by incumbent firms;³ and by encouraging entry of new, more efficient firms. Nickell (1996), Blundell et al. (1999) and Aghion et al. (2004, 2009), using firm-level data, show that product market competition has indeed a positive impact on productivity. However, some disagreement exists on the impact of competition – and of competition enhancing policies – on innovation. Whinston and Segal (2007) study a dynamic Schumpeterian model in which incumbents and innovating entrants compete for the market. They find that pro-competitive policies that facilitate entry tend to increase entrants' incentives to innovate by front-loading the returns from their R&D investments. Contrastingly, Acemoglu and Cao (2010) propose a model where both the incumbents' and entrants' innovation rates are endogenous, finding that subsidies to entrants may reduce productivity growth by curbing incumbents' innovation.⁴

The disagreement on the effect of competition – and hence competition enhancing policies

²More generally, competition acts as a selection process that reallocates market shares in favor of the most productive firms. Haskel (2000) provides empirical evidence of this process. Disney et al. (2003) and Syverson (2004) show that competition reduces productivity dispersion suggesting that inefficient firms are forced to either catch-up or to exit.

³Competition also presses managers to reduce x-inefficiency (Hicks, 1935, Leibstein, 1966). This point is made theoretically by Nalebuff and Stiglitz (1983), while Vickers (1995), Nickell et al. (1997), Griffith (2001) and Bloom and Van Reenen (2007) provide empirical evidence of a positive relationship between competition and x-efficiency.

⁴Bartelsman and Doms (2000) empirically show that a large fraction of productivity growth is indeed due to incumbents' incremental innovation. Aghion et al. (2009) study the effect of the threat of technologically advanced entry on incumbents' innovation. They find that such a threat increases innovation in sectors close to the technology frontier, where an innovating incumbent can survive the entry of technologically advanced firms. Yet, it reduces innovation in laggard sectors where the threat of technologically advanced entry decreases the incumbent's expected rent from innovating.

– and innovation is also reflected in the extensive macroeconomic literature on Schumpeterian growth. At first sight, the intuitive Schumpeterian argument that firms invest and innovate to capture future monopoly rents suggests a negative relationship between competition and innovation. However, this intuition, which is reflected in early Schumpeterian growth models like Aghion and Howitt (1992) and Caballero and Jaffe (1993), has been overturned by several recent contributions. In particular, Aghion et al. (2001) study a model of step-by-step innovation where both leaders and laggards produce and innovate. Laggards must first reach the leader’s technological level before being able to challenge its leadership and replace it. Aghion et al. (2001) find that in most cases an increase in competition spurs innovation, as the standard negative effect linked to lower rents is dominated by a positive ‘escape-competition-effect’. Aghion et al. (2005), which can be considered to be the benchmark in the literature on competition and innovation, further develop this approach taking into account the probability that an industry is in a neck-and-neck situation. They predict an inverted U shape for the relationship between competition and innovation, and find this prediction to be confirmed by firm-level data. A different link between competition and productivity is studied in Acemoglu et al. (2006), who examine the process of selection of more efficient firms and managers induced by competition. They show that this selection is beneficial for countries close to the technological frontier where its effect on cutting-edge innovation is more important. Yet, selection may be harmful for countries far away from the frontier, where the intensity of investment to adopt existing technologies is more important and is reduced by stronger competition.⁵

In light of the previous discussion, in principle we cannot rule out that competition policy, if too strict, may also have some adverse effects on efficiency. This identifies an issue that we will empirically address in section 3. However, it seems important to point out some reasons why the ambiguity of the impact of competition on innovation may not extend to competition policy. First, even if the relationship between *competition* and innovation has an inverted-U

⁵Denicolò and Zanchettin (2009) also analyze the role of competitive selection on growth in a model where competition is less extreme and innovations are not ‘drastic’. More efficient and less efficient firms coexist for several periods and the market share of more efficient firms gradually grows at the expense of the less efficient ones. They show that an increase in competition has conflicting effects on incentives to innovate: equilibrium prices go down, reducing profits from innovating, but the faster reallocation of market shares increases the innovators’ profits. The net effect on innovation, when competition is tough, is however positive.

shape, *competition policy* is less likely to have a strong impact in those markets where competition is already intense. Indeed, in most areas of competition law (i.e. vertical agreements, abuses of dominance and mergers) the pertinent prohibition applies only if the relevant market *significantly* departs from perfect competition (e.g. high concentration, high barriers to entry, large switching costs, etc.).⁶ As for cartels, even if the prohibition applies irrespective of the competitive conditions of the market, they generally represent the most serious restriction of competition. Moreover, the idea that cartels foster innovation has been generally dismissed (Nocke, 2007). Second, in most jurisdictions, all the relevant antitrust prohibitions (again with the exception of cartels) admit an 'efficiency defense'. This defense is meant to allow conducts that, although reducing competition, improve efficiency and benefit consumers. Therefore, the 'efficiency defense' provides a protection for the investments firms make to innovate.⁷

Our CPIs reflect the extent to which the various competition policy regimes allow this defense, and therefore incorporate the protection of investments in the interpretation of the antitrust rules. Hence, our measure of competition policy takes a higher value (*ceteris paribus*) where the protection of investments is a goal that shapes the interpretation of the applicable rules. Combining these two considerations, we should expect a positive relationship between *good* competition policy and innovation, both because competition policy increases competition only (or mostly) when the relevant market is in the first part of the inverted-U curve, and because competition policy refrains from increasing competition if this is likely to result in inefficiencies and/or less innovation.⁸

⁶In many of these areas antitrust law defines 'safe harbors' in terms of market shares or concentration indexes which establish a presumption of legality. For instance, in the European Union the legal and absolute presumptions are that some vertical restraints are compatible with competition law if none of the parties of the agreement has more than 30% of the relevant market.

⁷Baker (2007) argues that the application of modern economic theory has helped antitrust agencies to identify the types of firm's conduct and industry settings where antitrust interventions are most likely to foster innovation. Similarly, Gilbert (2008) maintains that antitrust policy has recognized the importance of finding a right balance between providing incentives to innovate and limiting practices that may harm competition.

⁸Of course, we are not claiming that our argument applies to any aspect of competition policy and to any antitrust decision. There may be specific interventions that may have a less positive effect on innovation and productivity growth. This may occur if a competition authority wrongfully believes that a high level of concentration is a sign of weak competition, while in fact it is the result of the selection process that characterizes intense competition. We believe that this type of error is less likely nowadays than it used to be 20 or 30 years ago. Indeed, the idea that the degree of concentration is a poor indicator of (the lack of) competition is now widespread in the daily work of many antitrust agencies around the world. Hence, our view is that these cases are likely to represent exceptions and therefore should not alter the positive relationship between competition policy and innovation.

Competition policy is embedded in a wider and interconnected system of institutions and policies that might present inherent complementarities (Aghion and Howitt, 2006). In our context, legal institutions stand out as particularly relevant, since the enforcement of competition law is intimately linked to the functioning of the judiciary system for several reasons. First, competition law is enforced by public bodies and by private firms and individuals who can bring suits in courts for alleged anti-competitive conducts. Second, in some jurisdictions the competition authority can only challenge a conduct or a merger before a court. Finally, even in those jurisdictions where the competition authority acts as an 'adjudicator', its decisions are subject to judicial review, so that courts have the last say on all competition policy interventions.

The interaction between a country's legal rules and economic activities has recently attracted a large interest following the path-breaking work by La Porta et al. (1997, 1998) who argue that legal traditions spread around through conquests and colonization and shaped the subsequent evolution of legal and regulatory institutions. It has been shown that legal origins affect many other dimensions including bank ownership (La Porta et al. 2002), entry regulations (Djankov et al. 2002), labor market regulation (Botero et al. 2004), and government ownership of the media (Djankov et al. 2003a). Some studies also looked at how the characteristics of the judiciary and other government institutions affect the security of property rights and contract enforcement (Djankov et al., 2003b; La Porta et al., 2008). On the basis of the results by Djankov et al. (2003a) and La Porta et al. (2004) we expect that a lower level of formalism of the judicial procedures and greater judicial independence should improve the quality of the judicial review of the decisions made by competition authorities. Hence, we expect positive complementarities between several indicators of the quality of the judiciary system and competition policy.⁹ In doing this, we are close to the recent work of Aghion and Howitt (2006), and more generally to the literature on institutions and long-term economic performance as surveyed in Acemoglu et al. (2005), Glaeser et al. (2004), and Beck and Levine (2005).

⁹Recently, Malmendier (2009) critically discusses the literature on the nexus between law, finance, and growth. Analyzing the role of the Roman shareholder company, she provides empirical support for the view that political institutions can dominate the role of legal institutions in shaping economic performance. She concludes by suggesting a cautious use of the legal origin approach to measure the transaction costs of institutional environment. The debate is still unsettled and it is not the aim of this paper to enter it.

More importantly, our paper contributes to the still very limited empirical literature that evaluates the effectiveness of competition policy. Dutz and Hairy (1999) and Dutz and Vagliasindi (2000) use a cross-section of 52 countries and a small sample of transition economies respectively and find a positive effect of antitrust effectiveness on GDP growth. However, they use 'subjective' measures of competition policy that are based on the perceptions of market participants which, as a consequence, may not correctly represent the objective features of a competition policy regime. Konings et al. (2001) and Kee and Hoekmann (2007) look at the impact of the introduction of competition policy on industrial mark-ups in two very different samples (the first one includes Belgium and the Netherlands and the second includes a large panel of industries in developed and developing countries). Neither paper finds direct evidence of a positive effect of the introduction of competition policy or competition law on mark-ups.¹⁰ However, the interpretation of the results might be misleading as the employed measure of competition policy appears inadequate to capture those features that are likely to impact on its effectiveness.

Finally, especially for the empirical approach, our work is closely related to the literature that examines the impact of regulation and other competition enhancing policies on productivity growth. Nicoletti and Scarpetta (2003) focus on the direct effect of privatization and liberalization on TFP growth. They show that market-oriented regulatory reforms significantly contributed to improving productivity in OECD countries during the Nineties, especially by reducing the gap to the technological frontier.¹¹ Pavcnik (2002) finds a direct impact of trade liberalization on productivity improvements that works through the reallocation of resources to more efficient producers. Several other papers, instead, look at the effect of competition and entry on productivity growth (e.g. Griffith and Harrison, 2004, and Aghion et al., 2009). They use policy variables, such as the introduction of the EU single market program or the UK privatization program, as instruments for competition, which is proxied by the price-cost

¹⁰See also Sproul (1993), who finds that prices increase in industries after a cartel has been discovered and convicted; Clarke and Evenett (2003), who find that the vitamin cartel reduces cartel prices in jurisdictions where antitrust conviction is more likely and costly; and Voigt (2009), who finds a positive effect of a set of indicators of the quality of competition policy on total factor productivity, that however disappears when controlling for institutional quality.

¹¹This results are partially critically challenged by Bourlès et al. (2010) and Amable et al. (2009).

margin, and entry. They show that the policies have a positive impact on competition and entry and these, in turn, increase productivity. Unlike these latter studies, we do not attempt to measure the channel through which competition policy affects productivity. First, this is not essential to our exercise as we want to assess the policy effectiveness. Second, in this way we avoid specifying any notion of competition which might be problematic both theoretically and empirically.¹²

3 Econometric Specification

To make robust causal inference on the effectiveness of competition policy, we analyze the direct link between the policy and TFP growth.¹³ Our empirical implementation builds on a general quality-laggard framework typical of endogenous growth models (e.g. Aghion and Howitt, 2006). The basic idea is that laggard industries/countries can catch up with the technological frontier by innovating or adopting the leading technologies. Therefore, the technological and organizational transfer from technology-frontier's firms influences the productivity of laggard industries and, hence, their productivity is co-integrated with that of the leader. Under the assumption of long-run homogeneity, this process has an Error Correction Model (ECM) representation where the industry-level TFP growth ($\Delta TFP_{i,j,t}$) in country i and time t depends on the technology transfer from the country on the technological frontier ($TFP_{L,j,t}$), and the productivity gap or distance to the technological frontier ($TFP_{L,j,t}/TFP_{i,j,t}$) (e.g. Griffith et al., 2004, pg. 886). These dimensions constitute sources of observed heterogeneity that should explain productivity growth and, hence, should be empirically controlled for.

Clearly, the rates of TFP growths are affected by other country-industry characteristics. From our previous discussion, competitive pressure is one of these important drivers. In particular, following the theoretical framework proposed by Aghion and Schankerman (2004) and

¹²For instance, from a theoretical point of view, the price cost margin (PCM) is a poor indicator as it (imperfectly) captures only a short-run notion of competition. Even in this case, the relationship can be non linear and an increase in competition may result in a higher PCM (Boone, 2000).

¹³While under strict neoclassical assumptions, TFP disembodies technical change or dynamic efficiency, in practice it integrates a range of other efficiency effects including those from organizational and institutional change, changes in returns to scale, and unmeasured inputs such as research and development and other intangible investments (e.g. Inklaar et al., 2008). Moreover, industry-level TFP also captures the effects of reallocation of market shares across firms.

Acemoglu et al (2006), and the empirical approach suggested by Nicoletti and Scarpetta (2003), Griffith and Harrison (2004), and Aghion et al. (2009), we assume that competition-enhancing policies – such as competition policy (*CPI*), product market regulations (*PMR*), as well as trade liberalization – are some of the main drivers of this residual heterogeneity which is not captured in the quality-ladder framework.

Moreover, following Griffith et al. (2004), we also assume that other observable industry-country-specific factors connected to innovation – such as R&D intensity (*R&D*) and human capital – directly affect the rate of TFP growth.¹⁴ Finally, following the existing literature (e.g. Nicoletti and Scarpetta, 2003 and Griffith et al., 2004) we model the remaining unobserved heterogeneity by means of an error term, which takes the form $\varepsilon_{i,j,t} = \psi_{i,j} + \phi_t + u_{i,j,t}$. The country-industry-specific fixed-effects $\psi_{i,j}$ account for the time-invariant unobserved heterogeneity and the full set of time dummies (ϕ_t) controls for common macroeconomic shocks that may affect TFP growth in all countries at the same time.¹⁵ The basic equation that we estimate is thus the following:

$$\Delta TFP_{i,j,t} = \alpha + \beta CPI_{i,t-1} + \delta \Delta TFP_{L,j,t} - \sigma \frac{TFP_{L,j,t}}{TFP_{i,j,t}} + \gamma X_{i,j,t-1} + \chi Z_{i,t-1} + \varepsilon_{i,j,t}, \quad (1)$$

where $CPI_{i,t}$ is one of our indicators of competition policy in country i at time t , $X_{i,j,t-1}$ are country-industry-specific control variables (human capital, trade openness, R&D, and a country-industry-specific trend), $Z_{i,t}$ are country-specific controls (product market regulation and the quality of institutions).¹⁶

As we mentioned in section 2, some recent papers (e.g Aghion et al., 2005, and Acemoglu et al. 2006) suggested that competition-enhancing policies may also influence TFP growth

¹⁴Differently from them, however, we do not analyze how R&D might indirectly affect TFP growth by shaping the catch-up process.

¹⁵We run a large amount of alternative specifications to analyze how these assumptions on the error terms affect our results. This discussion is reported in more details in appendix C. Neither the choice of different individual effects, nor the accounting of potential serial correlation in the residuals affects our main results.

¹⁶Potentially, competition policy might have a non-linear effect on productivity growth akin to the non-linear effect of competition on innovation found in the literature (Aghion et al., 2005). In section 2 we theoretically motivate why we do not think that such a non-linear effect should be observed. To empirically validate our claim, we tried two alternative specifications. First, we used a quadratic, rather than a linear, term for the Aggregate CPI. Second, we used a step function for low, medium, and high levels of the Aggregate CPI. In both cases we do not find evidence of such non-linear effect, which make us confident of the chosen specification (1).

through an indirect channel, by interacting with the distance to the technological frontier.¹⁷ Indeed, competition policy, by increasing competition and reducing entry barriers, may increase the opportunities and incentives for the adoption of leading technologies. However, the returns from increasing productivity and improving efficiency in order to escape competitive pressure might be higher for firms competing neck-and-neck with rivals that are close to the technological frontier. Hence, the effect of competition policy might differ, depending on the level of technological development of a country-industry. We therefore look at an additional specification where the effect of competition policy on TFP is interacted with the technology gap.

3.1 Identification

The identification of a causal link between competition policy and productivity growth crucially relies on the ability to account for the potential endogeneity of our key policy variables. Especially when looking at country-level aggregates, endogeneity might arise from omitted variable bias as well as from two-way causality and measurement errors. In this paper we adopt a multi-steps approach, using several alternative strategies to pursue the ultimate goal of establishing a robust causal relationship between competition policy and TFP growth.

First, we believe that two-way causality is not a major concern in our case. In principle, the application of competition policy might be focused on less competitive and productive markets, which in turn might lead to a negative correlation between the CPIs and the error term. However, our CPIs aggregate several institutional characteristics, which are unlikely to respond swiftly to changes in TFP growth rates. Institutions face inertia and slowly evolve over time quite independently of specific and short-run changes in market outcomes.¹⁸ Even those variables that represent some relevant enforcement features, such as the human and financial resources, depend on political decisions that generally take time to be put in practice. In any case, in order to reduce the potential bias resulting from two-way causality, we use lagged val-

¹⁷Similarly, some empirical studies recently analyzed the differential effect of product market regulation on productivity and innovation depending on the distance to the frontier (Nicoletti and Scarpetta, 2003, Amable et al., 2009, Bourlès et al., 2010).

¹⁸For instance, the introduction of leniency programs or the adoption of the EU competition law model in Eastern European countries are likely to be the consequence of the diffusion of some institutional innovations, rather than a response to inadequate short-run market performances.

ues of the policy variables with respect to our dependent variable. This is a standard approach that relies on the assumption that the lagged values of the policy are uncorrelated with the error terms of the estimated equation (e.g. Griffith et al., 2004 use this exclusion restriction to identify the causal effect of R&D on industry TFP growth).

The main identification issue in the context of our model is related to the existence of an omitted variable bias. The panel structure of our data-set allows us to control for time-invariant unobserved individual heterogeneity at the industry-country level through fixed-effects as well as for time fixed-effects. However, there still might be time-varying unobserved heterogeneity. In particular, this might derive from the existence of several other competition-enhancing policies or, in general, other policies correlated with competition policy that might affect TFP growth rates. In our basic specifications, we control for those we believe to be the most prominent policies affecting competition (product market regulation, liberalization, and privatization) and for trade openness. While we are confident that these controls should help mitigate the endogeneity problem, we nonetheless propose a twofold approach to provide further evidence on the causal nature of the link between competition policy and productivity growth.

First, we propose an instrumental variable estimation, which allows us to explicitly test whether endogeneity matters and to control for another source of potential inconsistency of OLS estimates: the existence of measurement errors. We use two very different sets of instruments. Following some recent contributions which find political variables to determine policy outcomes (e.g. Besley and Case, 2000; Duso and Roller, 2003; Duso and Seldeslachts, 2010), we use the government type and its ideological position on regulatory issues as a first set of instruments. An alternative set of instruments derives from a well-established practice in industrial organization (e.g. Hausman, 1997). This consists of using different aggregations of the potentially endogenous variables in other markets as an instrument for the same variables in the market of interest. While the formulation of competition policy in a given country is likely to be affected by the evolution of competition policy in neighboring countries, the latter should not correlate with the rate of TFP growth in the country of interest. This provides the exclusion restriction necessary for identification. The existence of a correlation among policies in different countries is supported by the observable common trends in the evolution of competition

policy during the last decades. These trends are possibly due to the leading policy-setting role taken by jurisdictions such as the US or the EU, after which the other jurisdictions' policies are modeled. Moreover, a vigorous international academic and policy debate established a general consensus about the most efficient policies to adopt in the field of competition laws, which surely also generate common trends in its evolution over time.¹⁹

Second, in addition to the IV estimation, we adopt a less formal approach to improve our identification strategy by looking at potential non-linear effects of competition policy on TFP growth. We search for situations where we expect competition policy to have a differential effect on productivity as compared to other omitted factors or policies. If we were to observe this kind of behavior in the data, this would enhance our confidence that the estimated nexus between the quality of a competition policy regime and TFP growth can be interpreted in a causal way. Although one can never fully rule out the possibility that some complex interactions of omitted shocks would drive the results, this would then seem unlikely. There are two dimensions of heterogeneity that we think are important in this respect. The first is related to country-specific characteristics. As discussed in section 2, we expect competition policy to be more effective in those countries where the quality of legal institutions is higher. In fact, national courts are strongly involved in the enforcement of competition policy, as they often retain the power to adjudicate antitrust cases either directly or in appeal. Yet, crucially for our argument, courts are not involved in the adoption of other productivity-enhancing policies (for instance, regulation, R&D subsidies or fiscal policy) or, at least, they are involved only indirectly. The second dimension of heterogeneity we look at is related to industry-specific characteristics. Our data encompass industries belonging both to the manufacturing and service sectors. We expect the former to be significantly more affected by competition policy. The reason is that services are in general subject to strong sector-specific product market regulations – such as price control, entry regulations, and state ownership – which, in these industries, play a more significant role

¹⁹The role of multinational cooperation for the discussion and adoption of best practices around the world increased over the years covered in our sample. Such cooperation, which took place within the OECD and other international organizations, was fostered by the creation of the International Competition Network (ICN). This informal forum was initiated by the US in 1995 with the aim of providing a platform for competition authorities from around the world to discuss the whole range of practical competition policy enforcement and policy issues. The main objective of the ICN is exactly to spread best practice and promote convergence.

in shaping the competitive environment and, hence, productivity outcomes than competition policy. This intuition is empirically supported by Nicoletti and Scarpetta (2003) who find that deregulation plays a significantly greater role in fostering productivity in services than in manufacturing sectors. This kind of regulation clashes with competition policy, and for this reason we expect ex-ante that competition policy will be less effective in those industries where the tightness of product market regulation is greater.²⁰

4 Data Sample and Descriptive Statistics

We estimate our model (1) on a sample of 22 industries in 12 countries over the period 1995-2005. The countries included in the study are: Canada, the Czech-Republic, France, Germany, Hungary, Italy, Japan, the Netherlands, Spain, Sweden, the UK, and the US.²¹ We use data both at the national level and at the industry level. National level data are used to measure the policy variables (competition policy, product market regulation) and the quality of institutions. The remaining variables are measured at the industry level, which belong both to the manufacturing and to the service sectors.²²

In the following sections we introduce the main variables that we use in our regressions. We begin by discussing our main explanatory variables, the competition policy indexes. We then move to the discussion of the TFP growth measure and the other explanatory variables. We conclude by introducing our instruments.

²⁰Clearly, other forms of regulation – e.g. health and safety regulations – might have an additional effect on productivity growth also in manufacturing industries. However, these regulations are inherently different from those policies that directly control the competitive process and, hence, should not affect our identification argument.

²¹These countries have been selected to be representative of different legal systems (common law and civil law), to include both EU and non-EU countries and, among the EU countries, both founding members and countries that have recently entered the Union, namely Hungary and the Czech Republic.

²²The 22 industries (ISIC rev.3 codes) included in the study are the following: agriculture, forestry and fishing; mining and quarrying; food products; textile, clothing and leather; wood products; paper, printing and publishing; petroleum and coal products; chemical products; rubber and plastics; non-metallic mineral products; metal products; machinery; electrical and optical equipment; transport equipment; furniture and miscellaneous manufacturing; electricity, gas and water; constructions; hotels and restaurants; transport & storage; communication; financial intermediation; business services.

4.1 Measuring the Quality of Competition Policy: The CPIs

The ultimate aim of competition policy is to maximize social welfare. Hence, the quality of a competition policy regime should be evaluated on the basis of the ability of this policy to deter firms that operate within its jurisdiction from undertaking those behaviors that, by impairing competition, reduce social welfare. In this section, we therefore provide a self-contained discussion on how we measure the quality of a competition policy regime. We shortly report on the theoretical background behind our data collection exercise, the measurement issues, as well as the steps of the aggregation process we undertook to generate a set of summary indicators of the quality of competition policy, the CPIs. An exhaustive discussion of all the issues touched upon in this section can be found in the companion paper (Buccirossi et al., 2010). Moreover, in appendix A we give a more in-depth overview of the properties of some of our indicators and their distributions.

Following Becker's (1968) theory of optimal deterrence, we consider that the level of deterrence is determined by three fundamental elements: the size of the sanctions, the probability of detection and conviction, and the probability of errors. Several institutional and enforcement features of a competition policy regime might affect these three factors (see Buccirossi et al., 2009). The features which we believe have the strongest impact on the level of deterrence of anti-competitive behaviors are: the degree of independence of the competition authority (or CA) with respect to political or economic interests (formal independence); the separation between the adjudicator and the prosecutor in a competition case (separation of powers); how close the rules that make the partition between legal and illegal conducts are to their effect on social welfare (the quality of the law on the books); the scope of the investigative powers the CA holds (powers during investigation); the level of the overall loss that can be imposed on firms and their employees if these are convicted (sanctions and damages); the toughness of a CA, which is given by its level of activity and the size of the sanctions that are imposed on firms and their employees in the event of a conviction, and the amount and the quality of the financial and human resources the CA can rely on when performing its tasks.

We collected information on each of these features, by asking several specific questions.²³

²³For instance, to measure the quality of the law, we collected information on the standard of proof that is required

We gathered these data separately for the three possible infringements of the antitrust legislation (hard-core cartels, other anti-competitive agreements, and abuses of dominance) and for the merger control policy in each country and for each of the years in the sample. Most of this information was directly obtained from the CAs of the 13 jurisdictions included in our sample through a tailored questionnaire.²⁴ The data obtained from this survey were integrated with information derived from the country studies carried out by the OECD in the context of its reviews of regulatory reforms, from the chapters on competition and economic performance in the OECD Economic Surveys and from the CAs' own websites and publications.²⁵ Despite this extensive data gathering exercise, we encountered some difficulties in obtaining data on the toughness of the CAs and we could include in our database only details on the maximum jail term imposed on managers of firms involved in hard-core cartels (for those jurisdiction that have this type of sanction) and the number of hard-core cartels and mergers investigated every year.²⁶

when deciding on a specific type of violation as well as the nature of the goals that inform the decision-making process. To measure the CA's powers during investigations we collected information on the power to impose, or request, interim measures; the powers to gather information by inspecting the premises of the firms under investigation or the private premises of the firms' employees; the powers to gather information by wiretapping the conversations of the firms' employees. Buccirossi et al (2010) describes all these issues in depth.

²⁴Our sample includes 12 countries and 13 jurisdictions, as it includes the European Union. We only surveyed the CAs which are either independent public bodies or ministerial agencies/departments, while we did not survey the courts (but we have collected data on their powers and activities). The bodies surveyed are: Competition Bureau (Canada); Urad pro ochranu hospodarske soutez (Czech Republic); Directorate General for Competition Affairs (European Union); Conseil de la Concurrence (France); Direction Gènèrale de la Concurrence (France); Bundeskartellaamt (Germany); Gazdasági Versenyhivatal (Hungary); Autorità Garante della Concorrenza e del Mercato (Italy); Japan Fair Trade Commission (Japan); Nederlandse Mededingingsautoriteit (Netherlands); Servicio de Defensa de la Competencia (Spain); Tribunal de Defensa de la Competencia (Spain); Konkurrensvetket (Sweden); Office of Fair trading (UK); Competition Commission (UK); Federal Trade Commission (US); Antitrust Division - Department of Justice (US).

²⁵Despite the active collaboration of most CAs, it was not possible to collect all data on the enforcement characteristics of the competition policy necessary to build the CPIs for the period considered. Hence, our database has some missing observations. We tried to fill the gaps by asking the CAs to provide us with an imputation of the missing observations based either on other data at their disposal or on their historical knowledge of the trends. When this was not possible, whenever this was allowed by the characteristics of the other available data on that specific feature, we performed some limited imputation of the missing data. Nevertheless, the database still has some gaps. This means that in some cases we do not have all the information necessary to calculate a specific index. To avoid calculating indexes whose value could be altered by the lack of information, we do not calculate an index (at any level of aggregation) if 50%, or more, of the relevant information content was missing.

²⁶It is therefore clear that our measure of enforcement is less accurate than our measure of institutions. However, our CPIs capture most of the features that have a likely impact on the deterrence properties of the analyzed competition policy regimes as they fully describe their institutional features and proxy the level of enforcement by important variables such as the budget dedicated to the implementation of this policy, the amount of human resources devoted to the same aim and their quality. Furthermore, we believe that the institutional features of a competition policy regime play the greatest role in determining its effectiveness. As Kovacic (2009, 145) recently pointed out "Good policy runs on an infrastructure of institutions, and broadband-quality policy cannot be deliv-

The CPIs have a pyramidal structure.²⁷ We collected data for each of the seven key feature of competition policy mentioned above. Each piece of information is then assigned a score, on a scale of 0-1, against a benchmark of generally agreed best practice (from worst to best).²⁸ The best practice is determined by relying on scientific papers and books, on documents prepared by international organizations such as the International Competition Network and the OECD, and on our judgement. All the information on a specific policy feature is summarized in a separate low-level index using a set of weights to linearly aggregate it.²⁹ We calculated separate indexes for each of the three possible competition law infringements and for mergers, to take into account the differences in the legal framework and, where possible, in the enforcement.³⁰

The low-level indicators are subsequently aggregated into two medium-level indexes for each of three types of possible competition law infringements and for mergers: one which summarizes the institutional features of the competition policy regime and one which summarizes its enforcement features. The medium-level indexes are then aggregated to form a number of different summary indexes. More specifically, we calculate (for each country and each year in the sample): i) one index that measures the deterrence properties of the competition policy regime with regard to all antitrust infringements (the Antitrust CPI) and one that measures its deterrence properties in the merger control process (the Mergers CPI); ii) one index that assesses the institutional features (the Institutional CPI) and one that assesses the enforcement features (the Enforcement CPI); iii) a single index that incorporates all the information on the competition policy regime in a jurisdiction (the Aggregate CPI).

The weights employed in this aggregation process are based on the relevance that each item, ered on dial-up-quality institutions." Hence, one can see good institutions as a necessary, yet possibly not sufficient, condition for a good enforcement.

²⁷Our methodology is akin to the one developed by the OECD for the indicators of product market regulations (PMR) and the competition law and policy indexes (CPL). See Boylaud, Nicoletti, and Scarpetta (2000), Conway and Nicoletti, (2005) Conway and Nicoletti (2006) for the former and Høj (2007) for the latter.

²⁸When a data entry is quantitative it is normalized by dividing it by the highest corresponding value held by any CAs in the sample, so that even quantitative information assumes a value between 0-1.

²⁹We are aware that there might be complementarities among different aspect of competition policy that we may miss by using this linearly additive specification. However, we believe that it would be difficult to choose a more precise approximation of the relationship that could exist between these variables. Hence, we have selected this aggregation form that has the advantage of being simple and at the same time rather complete.

³⁰This was not always easy. For example, the CAs rarely have separate divisions that deal with the different types of infringements, hence we could not obtain separate data on the resources employed for each of them. Hence, the resource index takes the same value for all the three possible antitrust infringements, as well as for merger control.

in our view, deserves.³¹ However, in order to check whether our choice of weights has a decisive influence on the results, we also use three alternative weighting schemes. The first uses an agnostic approach and weights each piece of information equally. The second, aggregates the features of competition policy using factor analysis.³² The correlation coefficients between the values of the Aggregate CPIs built with our weights and these two alternative CPIs built with equal weights and the weights obtained from the factor analysis take very high values (0.97 and 0.96 respectively) and they are significantly different from zero at the 1% level. In the robustness section we run our basic regression using the CPIs calculated by means of these alternative weighting schemes and show that the results are robust. The third alternative weighting scheme is based on random weights. We randomly generate, from a uniform distribution (0,1), 1,000 sets of weights, which are then normalized to sum to one. For each of these sets, we build one Aggregate CPI. In the results section, we report the distribution of the coefficients estimates for these 1,000 Aggregate CPIs and we show that our main findings are not affected.

4.2 Main Variables

In this section we describe the main variables that we employ in our regressions. We start by presenting TFP growth and then we move on to the control variables. All monetary measures are in real terms, using 2000 as the base year.

TFP growth. The dependent variable in our empirical model comes from the EU-KLEMS database.³³ TFP growth is measured by the Solow residual within the growth accounting framework as developed by Jorgenson et. al. (2005). Within this framework, TFP is measured under certain restrictive assumptions, among which that of prices equal to marginal costs. Following Griffith et al. (2006), we relax this assumption by multiplying the labor and capital

³¹We have been very conservative in the choice of the weights and we departed from equal weights only for situations for which there were robust theoretical reasons to do so. Moreover, we tried to be as transparent and explicit as possible in explaining why we chose each particular weight. The in-depth description of these issues can be found in Buccirossi et al. (2010).

³²A complete description of this alternative methodology and the results can be found in Buccirossi et al. (2010).

³³The EU-KLEMS project is funded by the European Commission, Research Directorate General as part of the 6th Framework Programme, Priority 8, 'Policy Support and Anticipating Scientific and Technological Needs'. The aim of the project is to create a database on measures of economic growth, productivity, employment creation, capital formation and technological change at the industry level for all European Union member states plus selected non-European countries from 1970 onwards. For a short overview of the methodology and results of the EU KLEMS database, see Timmer et al. (2007).

shares by the industry-level mark-up, which is estimated as the ratio between industry-level value added and labor and capital costs (see Paquet and Roubidoux, 2001).³⁴ In our sample, the average TFP growth at the industry level ranges between -1.7% for the business services sector and 3.7% for the communications sector. The average TFP growth in the entire sample is 0.0096%. A more in-depth description of this and other TFP-based variables can be found in appendix B.

Technology Gap. We use TFP levels to determine the technology frontier at the country-industry level and the technology gap between each country-industry and the frontier. Following the existing literature (Griffith et al., 2004; Nicoletti and Scarpetta, 2003), we obtain the technology gap using a two-step procedure. First, we calculate the ratio between the level of TFP in each country-industry and the geometric mean of the TFP levels in all the countries included in the sample for that industry. The frontier is defined as the country-industry with the highest ratio. Second, we obtain the technology gap by subtracting all the observed country-industry ratios from the frontier ratio.³⁵

R&D. The variable we use in our regressions is the ratio between R&D expenditure and the industry-level value added, both in nominal values. We gathered detailed data on the level of expenditure in R&D in different industries from the OECD Analytical Business Enterprise Research and Development (ANBERD) database, which covers 19 OECD countries, from 1987 to 2004. We took data on value added from the EU-KLEMS database. Unfortunately, data on R&D for the 'Agriculture, forestry and fishing' sector and the 'Mining and quarrying' sectors for all countries involved in the study as well as data for Hungary are not available in ANBERD.

Human Capital. We measure human capital as the share of high-skilled labor employed in each country-industry in a given year. We took data on human capital from the KLEMS database, which holds information on the level of educational attainment of workers by industry for all the EU member countries, the US and Japan from 1970 to 2004. Unfortunately, data

³⁴The concerns that we expressed on the ability of the mark-up to measure the intensity of competition in a market are not necessarily relevant for the correction implemented in the calculation of the Solow residual. Indeed, this correction cleans the TFP measure of the error due to the existence of a divergence between price and marginal cost (the mark-up).

³⁵Given the potential measurement errors in the construction of the Technology Gap (see appendix B), we test the robustness of our results using Labor Productivity (value added per worker) as a proxy for the distance from the technology frontier.

on Human Capital are not available for Canada.

Trade openness. We measure the degree of openness to trade by the ratio of industry import over value added in each specific industry. The data come from the OECD STAN database, which contains data on total exports and imports for 19 OECD countries, plus the EU, from 1987 to 2004, disaggregated by industry.

Product Market Regulation. We measure the tightness of product market regulation by the aggregate PMR index, taken from the OECD PMR database. The aggregate PMR index covers formal regulations in the following areas: state control of business enterprises, legal and administrative barriers to entrepreneurship, and barriers to international trade and investment. The tightness of regulation is measured at the national level on a scale between 0 and 6, where lower values indicate less tight regulation. Data on PMR are available for two years: 1998 and 2003.³⁶

Quality of Institutions. The quality of the institutions of a country enters in our regressions both as a control variable and as an interaction with the competition policy indexes in order to explore non-linearities in the effectiveness of competition policy. We use variables from four different sources to proxy the quality of the national institutions.

The first source of data is the World Bank Worldwide Governance Indicators (WGI) database, which collects aggregate and individual indicators for six dimensions of governance: voice and accountability, political stability and absence of violence, government effectiveness, regulatory quality, rule of law, control of corruption.³⁷ The data cover 212 countries and territories over the period 1996-2006 and are based on the views of a large number of enterprisers, citizens, and experts. We use the index that measures the national rule of law, as the most proper indicator of a country's legal system. The index takes values from -2.5 to 2.5, with higher values indicating better governance outcomes.

The second source of data is the Fraser Institute Database, which is used to construct the 'Economic Freedom of the World' indexes. From this database, we use an aggregate index (index_2) called 'legal system', which aggregates information on variables measuring judiciary

³⁶We assume regulation before 1998 to be as tight as in 1998, and regulation after 2003 to be as tight as in 2003. For the period between 1998 and 2003 we impute an average between the two available observations.

³⁷Note that all these indexes are highly correlated and contain, therefore, very similar information.

independence, impartiality of the courts, protection of intellectual property, law and order, and legal enforcement of contracts. These indexes, as the WGI, are based on the perceptions of enterprisers, citizens and experts. The indexes take values between 0 and 10, with higher values indicating better governance outcomes.

The third source of data is the Doing Business database of the World Bank and the International Finance Corporation, which collects data representing 'objective measures' of the overall quality of the regulatory and institutional environment on 181 countries. The data we use in our empirical model relate to the time and cost of enforcing debt contracts through the national courts system.³⁸ Finally, we use the legal origins dummies from La Porta et al. (1997).

Industry-level deviations from the trend. We use country-industry deviations from a linear and a quadratic trend to account for the effect of business cycles on TFP. When capacity is constrained, TFP growth may in fact reflect short-run demand fluctuations. We measure a different deviation from the trend for each country-industry using value added taken from the EU-KLEMS database.

4.3 Instruments for Policy

In our IV regressions we use two different sets of instruments for the policies (competition policy and PMR). First, we use political variables which are derived from the dataset developed by Cusack and Fuchs (2002) which uses two main sources:³⁹ the first is a database on political parties' programmatic position developed in the Manifesto dataset by Klingemann et al. (2006), while the second is the database developed by Woldendorp, Keman, and Budge (2000) on government compositions for 48 countries from 1948 onwards. For each country and year in our sample, we create measures of a government location along the Manifestos political dimensions by taking a weighted average of the programmatic positions of each of the parties belonging to government coalition. As weights, we used the number of each party's votes. We

³⁸The time of enforcing debt contracts represents the estimated duration, in calendar days, between the moment of issuance of judgment and the moment the landlord repossesses the property (for the eviction case) or the creditor obtains payment (for the check collection case). The cost of enforcing contracts represents the estimated cost as a percentage of the debt involved in the contract. For a full description, see Djankov et. al (2003b). Both variables have been measured within the Doing Business Project from 2004 on. In our specifications, we use the end of sample (2005) values, and assume it represents the quality of contracts enforcing for the entire sample period.

³⁹We are very grateful to Tom Cusack for providing us with the original data and the updates for the last years in our sample.

used the following programmatic positions:

Market regulation (per403). This variable measures favorable mentions in the parties' programs of the need for regulations to make private enterprises work better, actions against monopoly and trusts, in defence of consumer, and encouraging economic competition.

Economic planning (per404). This variable measures favorable mentions in the parties' programs of long-standing economic planning of a consultative or indicative nature.

Welfare state limitations planning (per505). This variable measures negative mentions in the parties' programs of the need to introduce, maintain or expand any social service or social security scheme.

European Community (per108): This variable measures favorable mentions in the parties' programs of the European Community in general, and on the desirability of expanding its competency.

Second, as we mentioned in the previous section, as additional instruments for the CPI and for regulation for a given country we use different aggregations of the level of these variables in other countries as possible instruments. In particular, we build different set of instruments based on country grouping (EU countries vs. non-EU countries). We then use as instruments for the policies (CPI and PMR) in one country the average value of these variables in all other countries from the same group, as well as the average value of these variables in all countries from other groups.⁴⁰

Table 1 reports the preliminary statistics for the main variables discussed in these sections.

4.4 Descriptive Analysis

As a first motivating step, we look at simple moments. We start by looking at the correlation between TFP growth and the CPI at the country-aggregate level. We compute a weighted average for TFP growth using the industry value added as a weight. The correlation coefficient is large and positive (0.29) and significantly different from zero at the 1% level. Figure 1 gives a graphical representation of this relationship at the country level. The positive correlation

⁴⁰Moreover, we also try using alternative instruments, such as the US policies as instruments for EU countries, the mean policies of EU member states (including the EC) as instruments for the US policies, and the mean between the EU and US policies for the policies in non-European countries such as Canada and Japan.

between the average TFP growth and the CPI is clear for most of the countries. In particular, we calculate a positive and significant correlation coefficient for the Czech Republic (0.83), France (0.32), Germany (0.43), Hungary (0.13), Japan (0.21), Netherlands (0.39), and UK (0.51).

Figure 1 also shows that there is substantial variation in TFP growth measures among the several industries within a country. In this study we also exploit this heterogeneity dimension, as competition policy might affect various industries in a different way. We make use of this argument as an additional step in our identification strategy. We therefore look at the pairwise correlation between the CPI and TFP growth at the industry-country level. Again, this correlation is positive (0.08) and significantly different from zero at the 1% level. Our empirical model starts from this simple correlation to identify the causal effect of the policy.

5 The Results

We start by considering the average effect of competition policy on total factor productivity growth by using the various CPI indexes discussed above. All regressions in the following tables include year dummies and industry-country fixed-effects. We further control for other competition-enhancing policies as measured by the OECD PMR index, trade liberalization, a country-industry-specific deviation from the trend to account for potentially different business cycles at the country-industry level, as well as for the other determinants of productivity growth, which we previously discussed. Most of the explanatory variables are lagged by one year to reduce possible endogeneity issues. Standard errors are clustered at the country level to allow for correlation among industries in the same country. We estimate the model by OLS. Our sample, after discarding some extreme outliers, consists of 1,847 country-industry-time observations.⁴¹

5.1 The Basic Model

In column 1 of table 2 we report the results of the basic specification. The key result is that the coefficient estimate for the Aggregate CPI is positive (0.0924) and statistically significant at the 1% level: good competition policy is strongly positively correlated to productivity growth in a

⁴¹We dropped the observations corresponding to the first and the last percentiles of the TFP growth distribution.

statistically significant way.⁴² This estimates also points to an economically significant effect. A coefficient estimate of 0.09 for the aggregate CPI implies an average elasticity of TFP growth with respect to the aggregate CPI of around 4.66.⁴³ Estimates for all other control variables conform to our expectations and to previous results reported in the literature and hence give us confidence about the quality of our specification. In particular, the TFP level of the leader, the technology gap, and import penetration have a positive and significant impact on TFP growth; while product market regulation, in the form of barriers to competition, has a negative effect on productivity growth, though this is not significant mimicking the findings by Nicoletti and Scarpetta (2003). Finally, the country-industry-specific trend that we inserted to account for short-run cyclical fluctuations in demand also has a positive and significant impact.

As we mentioned in section 4.2, there are two other important control variables – R&D and human capital – for which we unfortunately have many missing values.⁴⁴ Yet, we still want to analyze whether their introduction substantially affects our results, especially in light of potential omitted variable bias. In column 2 we therefore add R&D to our basic specification, which reduces the number of observations to 1,463. In line with Griffith et al. (2003), R&D intensity has a positive and significant impact on TFP growth. All other results, and especially the size and significance of the coefficient estimate for the Aggregate CPI, are not affected. In column 3, we report the results for our basic specification using the sub-sample where R&D is not missing. Again, our results are almost not affected. In column 4 we add to our basic specification human capital as a further control, which reduces the observation to 1,783. Again, this variable has a positive effect on TFP growth which, however, is not statistically significant. The other results are not substantially changed. We finally introduce both R&D intensity and human capital (column 5) and run our basic regression without these controls in the sub-sample where both variables are non-missing (column 6). Again, our main results are not affected, yet

⁴²This value is quite close to the value of the simple correlation coefficient that is equal to 0.08.

⁴³To give a more concrete idea of the economic meaning of this, we can look at one example such as the ‘food products’ industry in the UK. Over the period 2001-2004, the average productivity growth rate in this industry was 2.23%. Our model implies that part of this growth rate is due to the effect of the improvement of competition policy. In the same period, the average growth rate of the aggregate CPI was 3.75%. Using our average coefficient would imply that, had competition policy not improved, the average TFP growth rate would have been 1.92%.

⁴⁴In particular, R&D data are missing for Hungary and for several industries-years in other countries, while Human Capital is missing for Canada.

now the two controls are significant. This can be due to the sample selection effect, given that we run this specification on a much smaller sub-sample (1,408 observations). From this point on, we therefore decide to use our basic specification, so that we can use the maximum possible number of observations.⁴⁵

The last column (7) reports the results from the specification where we assume that competition policy might affect TFP growth differently depending on the country-industry's distance from the frontier. We therefore define three categories for the technology gap (low, medium, high) and allow the coefficient for the CPI to differ among them.⁴⁶ The estimated effect of competition policy is much larger and more significant (0.124) for country-industries far away from the frontier than for country-industries close to the frontier (0.053). This result is in line with the empirical findings of Nicoletti and Scarpetta (2003) who show that liberalization is mostly beneficial for productivity in manufacturing industries the further a given country is from the technology leader. Hence, increasing competition through an effective competition policy (or reducing entry-limiting regulations) may facilitate the adoption and development of advanced technologies, which increase productivity. The benefits of increasing competition in country-industries close to the technological frontier seem, instead, to be more modest, yet still positive and significant.⁴⁷

We then move to analyze the impact of the various dimensions of competition policy as measured by our disaggregated indexes. In table 3, we focus on the difference between institutions and enforcement in columns 1 and 2 and between mergers and antitrust in columns 3 and 4. Again, we obtain similar results to our basic model: the various dimensions of compe-

⁴⁵We do however run all regressions and robustness checks also adding R&D intensity and human capital as additional controls. These results can be obtained from the authors upon request.

⁴⁶We define the three dummies according to the distribution of the gap variable: low level (up to the 33rd percentile of the distribution), medium level (from the 33rd to the 66th percentile), and high level (from the 66th percentile).

⁴⁷These empirical findings might, at first glance, appear at odds with the theoretical framework proposed by Acemoglu et al. (2006), who show that a limited level of competition might be beneficial for sectors far away from the frontier, as we discussed in section 2. These are adopters and find it optimal to pursue an investment-based strategy rather than selecting high-skill managers and firms through a highly competitive process, which is necessary for innovation. Yet, our results do not necessarily refute this theoretical argument, as they might rather be driven by the fact that the country-industries in our sample are not, on average, so far from the technological frontier to switch to the investment-based strategy. This seems plausible in our context, as all countries in our sample are quite homogenous, being part of the OECD. Indeed, the empirical evidence put forward by Acemoglu et al. (2006) is based on data for non-OECD countries so as to approximate real technology 'followers', which are significantly behind the world frontier.

tion policy have a positive and significant effect on productivity growth. With the exception of the Antitrust CPI, the size of the effect is, however, always smaller than the one measured by the Aggregate CPI and, in some cases, it is also less significant. In particular, the results for the Enforcement CPI are the weakest, as the coefficient estimate drops to 0.04 and loses significance. Our interpretation for this result lies in the quality of the information summarized in this index. As we mentioned, we do not have complete measures of antitrust enforcement in terms of actions taken by the authorities but rather measures of the monetary and human capital resources.

The established positive and significant relationship between the quality of competition policy, and in particular of its institutional design in the area of antitrust, and productivity growth is the key finding of this study. As we discussed thoroughly in section 3.1, one major concern for the causal interpretation of this effect is the potential endogeneity of the policy. In this section we started tackling this issue by lagging the policy variables and controlling for most of the determinants of TFP growth discussed in the literature. The next sections aim at providing further evidence to get more confidence in the causal interpretation of the established link between competition policy and TFP growth.

5.2 Instrumental Variables

The next step that we propose in terms of identification strategy is to use an instrumental variables (IV) approach. The results of these IV estimations are reported in table 4. In the first three specifications (columns 1, 2, and 3), we use the political variables discussed in section 4.3 as instruments for the policy. Independent of whether we instrument only for the Aggregate CPI (column 1), for both the Aggregate CPI and PMR (column 3), or if we control for R&D while instrumenting both policies (column 2), we always find a positive and significant coefficient estimate for the Aggregate CPI, which is even larger than those reported in our basic OLS specifications. This result is reassuring, as IV estimates are consistent in the presence of endogeneity. The instruments used seem to work properly: they are correlated to the instrumented variables as shown by the high values taken by the F-statistic for the excluded instruments in the first-stage regressions. Furthermore, they are not correlated with the error term as shown

by the Sargan statistic.⁴⁸ Although being always consistent, IV estimates are not efficient in the absence of endogeneity. We therefore run a Wu-Hausman test of endogeneity and cannot reject the null hypothesis that the policies are exogenous at the 1% level, hence OLS estimates should be preferred because they are more efficient.

Even though, as we motivated, the proposed instruments seem to be a reasonable choice, one could still be concerned that they might be potentially correlated with other omitted factors. We therefore present a second set of results, based on a very different set of instruments. Following an established literature in industrial organization, we use the policies in neighboring jurisdictions as instruments for the policies in a given country. We instrument for the Aggregate CPI alone (column 4), for both the Aggregate CPI and PMR (column 6) and also control for R&D while instrumenting for both policies (column 5). Again, we consistently estimate a positive and mostly significant coefficient for competition policy. Similarly to the previous specifications, the instruments seem to be good in terms of correlation to the potentially endogenous variables (F-statistic for the excluded instruments), while they are uncorrelated to the error terms (Sargan test).⁴⁹ Moreover, also in this case the Wu-Hausman test cannot reject the null hypothesis of exogeneity, which might also partially explain the reduction in the significance level, as the IV estimates are less efficient than OLS estimates.

These sets of results confirm our claim that the established positive link between competition policy and productivity growth can be interpreted in a causal way, as we can reject the hypothesis that the policies are endogenous. Therefore, from now on we will focus on the OLS estimates which, in the absence of endogeneity, are more efficient.

⁴⁸In table 5 (columns 1 to 3), we report the first-stage regressions for the IV specifications 1 and 3 of table 4. As expected, a pro-regulation attitude of the government (**per403**) and a pro-welfare limitation programmatic position (**per404**) are, respectively, negatively and positively correlated to the CPI and positively and negatively correlated to PMR. A pro EU attitude (**per104**) correlates positively with the CPI and negatively with the PMR index, which is consistent with the tendency of the European Commission to support the development of more competitive markets.

⁴⁹In table 5 (columns 3 to 6) we report the first-stage regressions for the IV specifications 4 and 6 of table 4. The instruments are the mean of the policies in other countries from the same group (CPI.G and PMR.G) and a different group (CPI.NG and PMR.NG). While we could potentially expect a positive correlation if *all* policies move in the same direction, it is not a priori clear whether this should be expected for the mean policies over the entire sample period. Indeed, we report negative and significant average correlations.

5.3 Non-Linearities

The final, informal, step of our identification strategy is based on the exploitation of non-linearities. The idea is that competition policy is more effective in some countries than in others, due to their better institutional environment, and in those sectors which are less subject to industry-specific regulations. This should not be the case for other (omitted) policies. Moreover, the analysis of such non-linearities with respect to institutional details is an important contribution on a more theoretical basis, as it allows us to identify the existence of complementarities between competition policy and the efficiency of (legal) institutions and therefore to provide a novel contribution to a recently expanding literature (Aghion et Howitt, 2006). These results are reported in table 6.⁵⁰

In the first column, we present our basic specification where we simultaneously control for several institutional dimensions. Institutions seem to have a significant direct impact on productivity growth. Yet, unlike previous studies (e.g. Voigt, 2009), the positive and significant effect of competition policy is not affected by these additional controls. This reinforces the view that our indicators are able to capture the specific features of a competition policy regime, which we aimed to measure, and not the general quality of a country institutional environment.

In column 2 we then interact the Aggregate CPI with the dummies for legal origins. While the effectiveness of competition policy is significantly higher in countries with German and Nordic legal origins, it is clearly less so in countries with French legal origins, which in our sample are France, Italy, and Spain. These results seem to be in line with findings reviewed by La Porta et al. (2008) who report that countries with civil law are associated with a heavier-hand regulation, which has an adverse impact on markets and economic performance.

We then explore what specific characteristics of a legal system are important drivers of competition policy effectiveness. To exploit in the best possible way the limited variation in our institutional data and, at the same time, to allow for non-linear effects through a step function, we have transformed our continuous institutional variables into categorical variables based on their distribution. Thus, for each institutional variable we defined three dummies: low level

⁵⁰Notice that, for lack of space, we do not report the coefficient estimates for all control variables as they are anyway very similar to those reported in our previous regressions.

'*l*' (up to the 33rd percentile of the distribution), medium level '*m*' (from the 33rd to the 66th percentile), and high level '*h*' (from the 66th percentile) of institutional quality. Finally, we interact these dummies with the Aggregate CPI.

In column 3 we report results for the specification where we interact the Aggregate CPI with dummies measuring the cost of enforcing contract (EC).⁵¹ Although competition policy seems to have a positive and significant effect independently of the levels of contract enforcement, the effect is substantially larger – indeed more than double (0.240) – for those countries with low enforcement costs (CPI|IEC). Hence, our results support the view that competition policy effectiveness might be reinforced in countries where law enforcement is more efficient. In columns 4 and 5 we report the results of the specifications where we interact the Aggregate CPI with the Fraser 'Rule of Law' (RL) index and the WGI's 'Legal System' (LS) index.⁵² In both cases, we observe competition policy to be less effective in countries with less efficient legal institutions, such as a low rule of law or a poor legal system.

The reported results point out to complementarities between competition policy and some dimensions of legal institutions. This does not mean that policies in countries with a worse legal system or higher costs of enforcing contracts must be ineffective, but rather that their (partial) ineffectiveness can be better explained by the bad functioning of the more general legal institutions. Therefore, policy changes in this country must be adequately designed to account for the additional constraints imposed by the legal system.

The second dimension of heterogeneity of the degree of competition policy's effectiveness is industry-specific. As we pointed out, most of the service industries in our sample (e.g. electricity, gas, water, communication, financial intermediation) are subject to more or less heavy-handed sector-specific regulations and the organization of competition matters in these industries is delegated to sectoral authorities. Our claim is therefore that competition policy should have less of a bite in such industries, but this should not necessarily be true for other productivity-enhancing policies (e.g. fiscal policy and labor regulations). We report the results

⁵¹Very similar results are obtained by using the general index for contract enforcement. However, in that case we lose Italy since there is no information on the time needed to enforce the contracts for this country.

⁵²We also try specifications where we use sub-components of the legal system index, specifically 'Independence of the Judiciary' and 'Impartiality of the Courts' and find similar results.

of the specification where we estimate separate coefficients for the Aggregate CPI as well as for PMR in service and manufacturing sectors in column 6 of table 6. For the Aggregate CPI, we find a large (0.143) and statistically significant coefficient estimate in manufacturing, while the coefficient is much smaller and not significant in the service industries. Moreover, similarly to Nicoletti and Scarpetta (2003), we also find that the coefficient of product market regulation is negative and significant in services but not in manufacturing industries.⁵³ These results perfectly conform with our expectations.

All results reported in this section point to the existence of significant and sizable non-linear effects of competition policy on productivity growth. The estimated differential effects should not be expected for other kinds of policies, which might constitute our problematic omitted factors and generate endogeneity issues that would invalidate our causal inference. Hence, these further results might be seen as an additional step, which makes us more confident of the causal nature of the link we identify.

5.4 Extensions and Robustness Checks

We finally perform several robustness checks by using different CPIs and different measures for productivity growth, as well as different sample sizes.

First, to show that our results are not driven by the subjective weights we have chosen to build the CPIs, we use the three alternative weighting schemes, which were discussed in brief in section 4.1. In column 1 and 2 of table 7, we report the results obtained when using the Aggregate CPI constructed using equal weights or the weights generated by factor analysis, respectively. Our qualitative results are unchanged and competition policy still has a positive and significant impact on TFP growth at the 1% and 5% level, with a point estimates for the policy effect of 0.0925 and 0.0726, respectively. As an additional robustness check, we run 1,000 regressions, each using a different Aggregate CPI generated with a different set of weights randomly drawn from a uniform distribution (0,1). We therefore obtain estimates for 1,000 β coefficients and their relative t-statistics, whose distributions are represented in figure 2. The

⁵³We also tried to disaggregate this result even more and estimate industry-specific coefficients for the Aggregate CPI and the PMR indicators. The Aggregate CPI has a significant impact exclusively in manufacturing industries while the PMR indicator mostly in service industries.

distribution of the coefficients, which is represented in the first panel, ranges between 0.052 and 0.11, with a mean value of 0.084, which is close to our estimate in the basic specification. As shown by the second panel of the figure, all of the 1,000 coefficient estimates are statistically significantly different from zero (the lowest t-value is 2.98).

A second concern with the CPIs relates to the role of the EU competition policy in the EU member states. To correctly evaluate the effectiveness of each EU member state's competition policy, it is necessary to account for the fact the EU competition policy works alongside the national one. Therefore, for these countries, we have built a set of CPIs which are an average of each member states individual index and the EU index.⁵⁴ The coefficient estimate for the Aggregate CPI is still positive, highly significant and larger in size (0.115) with respect to our basic specification. This means that EU competition policy improves, on average, the effectiveness of national competition policies.

Third, we need to consider the limitations of the TFP measure we use. Until now, following Griffith et al. (2004), we have used a measure for TFP growth corrected for the mark-ups (as measured by the PCM) to account for imperfect competition. However, one may have some concerns about the quality of an industry-level aggregated PCM measure. Hence, we propose an alternative specification where we use TFP measures (i.e. the growth rate, TFP of the leader, and the technology gap) which are not corrected for the mark-ups. The coefficient estimate reported in column 3 is still positive and significant at the 10% level.

Fourth, while TFP growth is constructed using detailed information on labor and capital input (see appendix B) provided by the KLEMS, the Technology Gap uses OECD data, which are provided at a less detailed level of aggregation.⁵⁵ For this reason, we employed as an alternative a much simpler measure of productivity to measure the technology gap: labor productivity, as measured by value added per worker. In this specification, we kept TFP growth as our dependent variable and used TFP growth on the frontier as an independent variable (though the

⁵⁴Unfortunately, DG Competition did not provide us with information on enforcement features (such as the budget and the composition of the staff), at the EU level. Hence, we can only use information about EU institutional features. The precise definition of the variable is thus as follows: $AggregateCPI_{EU_{it}} = \frac{2}{3}(0.5 * Institutions_{CPI_{it}} + 0.5 * Institutions_{CPI_{EU,t}}) + \frac{1}{3}Enforcement_{CPI_{it}}$

⁵⁵Unfortunately, we could not employ the KLEMS data to construct the technology gap, since the KLEMS does not publish the series on capital stock and labor for all countries with the necessary level of detail.

frontier is defined in terms of labor productivity). The coefficient estimate reported in column 4 is still positive and significant at the 1% level.

Fifth, one might be concerned with the frequency of the data. TFP measures change quickly over time as a response to demand shocks, while our policy measures, although showing some significant time variation, present much more inertia. We therefore change the frequency of the data and look at long-run effects. We propose three different specifications along this dimension. In the first one, whose results are reported in column 5, we take longer three-year lags for all explanatory variables. Still, the coefficient of interest is similar in size to that of our basic specification, though it loses a bit of significance, as expected given the long lag used. In the second robustness check (column 6), we define TFP growth over a time span of three years, and sum up the figures from year t to year $t + 2$. We then 'lag' all explanatory variables by taking their value at the initial year, i.e. we look at how the value of competition policy in year t affects TFP growth between year t and $t+2$. In doing so, the number of observations is obviously reduced. We still find a positive and significant coefficient estimate (0.332) for the Aggregate CPI. As expected the coefficient is much larger, as it represents the effect of the policy on the three-year TFP growth rate. In the final specification, we use three-year averages for all variables (column 7). Also in this case, the coefficient estimate for the Aggregate CPIs is positive (0.0903) and significant.⁵⁶

Sixth, one might be concerned that the right level of aggregation of our data should be the country rather than the industry, as the main interest of our study is in the impact of a national policy. In Section 4.4 we reported a significant simple positive correlation between country-level TFP growth and competition policy. In this robustness check, we re-estimate our model by taking weighted averages of all our industry-specific variables using the value added of the industry as a weight (column 8). Also in this case, the coefficient estimate for the Aggregate CPIs is positive (0.0417) and significantly different from zero at the 10% level.

Finally, given the heterogeneity of competition policy's effectiveness across countries and industries, one might be concerned that our average results do not hold to the exclusions of

⁵⁶Similar, though a bit less significant, results are obtained using a five-year interval. The loss of significance is due to the imprecision of the point estimation deriving from the reduction of the data variability via the aggregation process.

particular countries and/or industries. We therefore run our basic regression on several sub-samples, sequentially excluding one or two countries (156 sub-samples) or one or two industries (506 sub-samples). For each sub-sample, we run our basic regression. The distribution of the β coefficients and their t-statistics are represented in figures 3 and 4. In all sub-samples, our estimates for the CPI are positive and, in the very large majority of the cases (99.4%), they are statistically significant at the 10% confidence level at least. While none of the estimates are insignificant when we exclude one or two industries, only in 4 out of the 156 sub-samples where we *simultaneously exclude two countries* are the coefficients significantly positive (one-tailed test) yet not significantly different from zero (two-tailed test).⁵⁷

5.5 Where does Identification comes from?

In appendix A we show that there is significant and quite continuous within-country variation in the Aggregate CPI in almost all countries, which identifies our policy effect. Nevertheless, in this section we try to spot which specific policy changes in the Aggregate CPI might be the major identifier of the average increase in TFP growth estimated in our regressions.

In figure 5 we plot the evolution of the average *residual* TFP growth and its 95% confidence interval across the 22 industries of each country as well as the competition policy indexes over the period 1995-2005. To mimic our estimation and control for sources of observable heterogeneity, we use the *residual* component of TFP growth which is not explained by the fixed-effects and the other variables included in our model (1) – excluding, of course, the Aggregate CPI. Again, we observe clear correlation patterns between the evolution of the Aggregate CPIs and of the residual average TFP growth. Our attention focuses on the evolution in the subset of countries and time periods for which the changes in policy are more noticeable and, therefore, which are most likely to influence the average effect identified in our estimation.

The first country that appears to drive the estimated relationship is the Netherlands: the residual TFP growth rises toward the end of our sample period, and then decreases between

⁵⁷The only specification for which the t-value is further apart from a critical level (p-value of 0.21) is when we *simultaneously* exclude the UK and the Czech Republic. The reason is that the coefficient estimates drops to 0.04, while the standard error increases a bit with respect to our basic specification. Notice, however that, even in this unique case, we still cannot reject the null hypothesis of the coefficient being positive at the 10% significance level with a one-tailed test.

2004 and 2005. The same evolution is associated to the aggregate competition policy index, which rises in 2003 following an upward trend in the investment in human and financial resources and then goes slightly down, again because of a contraction in the resources allocated to the competition authorities. In the UK, over the 2000-2003 period, we also observe a strong correlation between the rise of residual TFP growth and the evolution of the aggregate CPI index. Such evolution is due to a steady growth in the financial and human resources available to the two CAs after the introduction of the Competition Act in 2000. In the USA, the period between 1999 and 2003 seems to be the one that identifies a positive link between residual TFP growth and the CPIs, as the two series follow a much correlated pattern. The residual productivity growth performance is accompanied by an increase in the budget/gdp ratio in the US competition authorities, as well by an increase in the human resources.

In Hungary, we observe a common upward trend in residual productivity growth and competition policy. The major institutional changes that mark the evolution of the Hungarian competition policy are the attribution of more investigative powers to the competition authority and the modification of the criteria to sanction firms. The latter are no longer based on discretionary decisions of the competition authority, but are based on firms' turnover. These new tools were introduced starting from 2000. Moreover, a budget increase took place in 2002. A similar common upward trend can be observed in the Czech Republic. Indeed, while the residual productivity growth is constantly increasing, the competition policy experiences a slight increase due to the larger amount of resources available to the competition authority. From the institutional side, an important change that happened around 1998 is the attribution to the competition authority of the power to investigate business' premises.

6 Conclusions

The aim of competition policy is to ensure that firms undertake the least possible number of behaviors that reduce social welfare by impairing competition. Hence, an effective competition policy is one that deters most anti-competitive practices. Since by deterring anti-competitive practices competition policy should make markets work effectively and foster efficiency, in this paper we evaluate the direct impact of competition policy on efficiency. Hence, we estimate the

effect of the key institutional and enforcement features of a competition policy, summarized in a set of indicators, the CPIs, on total factor productivity growth in 22 industries of 12 OECD countries between 1995 and 2005.

Our results imply that good competition policy has a strong impact on TFP growth. The coefficient for the Aggregate CPIs is positive and statistically significant in a variety of specifications of our model. The Aggregate CPI also remains highly significant when we control for R&D, human capital, and the quality of a country's institutions. All these variables have a direct impact on TFP growth but do not alter the fact that competition policy is effective in increasing productivity. We obtain similar results when we look at a more disaggregated picture and separately consider the effects of a competition policy's institutional and enforcement characteristics and when we differentiate between the policing of antitrust infringements and the merger control discipline. Yet, the institutional and the antitrust elements of the competition policy appear to have the strongest impact on TFP growth. We adopt a multi-steps approach to identification based on instrumental variable regressions and the exploitation of non-linearities. We therefore provide careful support to the causal nature of the established link between competition policy and TFP growth. Furthermore, we observe complementarities between competition policy and the quality of legal institutions. The effect of the former is indeed larger in those countries where the enforcement costs are low and with a better legal system. Finally, our main findings prove to be robust to several checks, such as various measures of productivity, different aggregation techniques for the CPIs, and several sub-samples.

Our results provide support for the argument that competition policy creates gross benefits to the long-term performance of a country's economy. Nevertheless, these benefits should be compared to the costs of introducing competition laws and enforcing competition policy to perform a clear welfare assessment. Unfortunately, we did not have access to sufficiently precise and encompassing cost estimates to allow us to undertake such an analysis, which could, however, be undertaken in future work subject to further data collection. There is also scope for further refinements. Currently, we have used data on 22 industries in 12 OECD countries over ten years, but it would be interesting to expand the database so as to include more countries over a longer time period and, particularly, to analyze the impact of the policy in less

developed economies, which are further apart from the technological frontier. Moreover, the CPIs could be improved by including more detailed information on the enforcement features, in particular on the sanctions that are effectively imposed on convicted firms and individuals and on the resources employed and the number of cases investigated by the EU Commission. However, such a refinement of the CPIs is difficult because of the lack of available data. Indeed, if competition authorities were to increase their accountability by collecting and keeping reliable data on the enforcement of competition policy in an easily accessible format, studying the effectiveness of competition policy would become much easier.

References

- Acemoglu, Daron and Dan Cao**, "Innovation by Entrants and Incumbents," mimeo 2010.
- , **Philippe Aghion, and Fabrizio Zilibotti**, "Distance to Frontier, Selection, and Economic Growth," *Journal of the European Economic Association*, 2006, 4, 37–74.
- , **Simon Johnson, and James A. Robinson**, "Institutions as the Fundamental Cause of Long-Run Growth," in Philippe Aghion and Stephen Durlauf, eds., *Handbook of Economic Growth*, Elsevier, 2005, pp. 1114–1174.
- Aghion, Philippe and Mark Schankerman**, "On the Welfare Effects and Political Economy of Competition-Enhancing Policies," *Economic Journal*, 2004, 114, 800–824.
- and **Peter Howitt**, "A Model of Growth through Creative Destruction," *Econometrica*, 1992, 60, 323–351.
- and —, "Appropriate Growth Policy: A Unifying Framework," *Econometrica*, 2006, 4, 269314.
- , **Christopher Harris, Peter Howitt, and John Vickers**, "Competition, Imitation and Growth with Step-by-Step Innovation," *Review of Economic Studies*, 2001, 68, 467492.
- , **Nick Bloom, Richard Blundell, Rachel Griffith, and Peter Howitt**, "Competition and Innovation: An Inverted-U Relationship," *Quarterly Journal of Economics*, 2005, 120, 701–728.
- , **Richard Blundell, Rachel Griffith, Peter Howitt, and Susanne Prantl**, "Entry and Productivity Growth: Evidence from Microlevel Panel Data," *Journal of the European Economic Association*, 2004, 2, 265–276.
- , —, —, —, —, and —, "The Effects of Entry on Incumbent Innovation and Productivity," *Review of Economics and Statistics*, 2009, 91, 20–32.
- Amable, Bruno, Lilas Demmou, and Ivan Ledezma**, "Product Market Regulation, Innovation, and Distance to Frontier," *Industrial and Corporate Change*, 2009, 19, 117–159.

- Arellano, Manuel and Steve Bond**, "Some Tests of Specification for Panel Data: Monte Carlo Evidence and an Application to Employment Equations," *Review of Economic Studies*, 2001, 58, 277-297.
- Baker, Jonathan B.**, "The Case for Antitrust Enforcement," *Journal of Economic Perspectives*, 2003, 17, 2750.
- , "Beyond Schumpeter vs. Arrow: How Antitrust Fosters Innovation," *Antitrust L.J.*, 2007, 74, 575.
- Baltagi, Badi H., Seuck H. Song, and Byoung Cheol Jung**, "The Unbalanced Nested Error Component Regression Model," *Journal of Econometrics*, 2001, 101, 357-381.
- Bartelsman, Eric J. and Mark E. Doms**, "Understanding Productivity: Lessons from longitudinal microdata," *Journal of Economic Literature*, 2000, 38, 569–594.
- Beck, Thorsten and Ross Levine**, "Legal Institutions and Financial Development," in Claude Ménard and M. Shirley, eds., *Handbook of New Institutional Economics*, Springer, 2005, p. 251-278.
- Becker, Gary**, "Crime and Punishment: an economic approach," *Journal of Political Economy*, 1968, 76, 169–217.
- Besley, Timothy and Anne Case**, "Unnatural Experiments? Estimating the Incidence of Endogenous Policies," *The Economic Journal*, 2000, 110, 672–694.
- Bloom, Nick and John Van Reenen**, "Measuring and Explaining Management Practices Across Firms and Countries," *The Quarterly Journal of Economics*, 2007, 122, 1351–1408.
- Blundell, Richard, Rachel Griffith, and John van Reenen**, "Market Share, Market Value and Innovation in a Panel of British Manufacturing Firms," *Review of Economic Studies*, 1999, 66, 529–554.
- Boone, Jan**, "Competitive Pressure: The Effects on Investments in Product and Process Innovation," *RAND Journal of Economics*, 2000, 31, 549–569.

– , “A New Way to Measure Competition,” *Economic Journal*, 2008, 118, 1245–1261.

Bourlès, Renaud, Gilbert Cette, Jimmy Lopez, Jacques Mairesse, and Giuseppe Nicoletti, “Do Product Market Regulations in Upstream Sectors Curb Productivity Growth?: Panel Data Evidence for OECD Countries,” NBER Working Paper No. 16520 2010.

Boylaud, Oliver, Giuseppe Nicoletti, and Stefano Scarpetta, “Summary Indicators of Product Market Regulation with an Extension to Employment Protection Legislation,” OECD Economics Department Working Paper, No. 226 2000.

Buccirossi, Paolo, Lorenzo Ciari, Tomaso Duso, Giancarlo Spagnolo, and Cristiana Vitale, “Deterrence in Competition Law,” WZB Discussion Paper SP II 2009-14 2009.

– , – , – , – , – , and – , “Measuring the Deterrence Effect of Competition Policy: The Competition Policy Indexes,” *Journal of Competition Law and Economics*, 2010, forthcoming.

Caballero, Ricardo J. and Adam B. Jaffee, “How High are the Giant’s Shoulders? An Empirical Assessment of Knowledge Spillovers and Creative Destruction in a model of economic Growth,” *NBER Macroeconomic Annual*, 1993, 8, 15–86.

Caves, Douglas W., Laurits R. Christensen, and W. Erwin Diewert, “The Economic Theory of Index Numbers and the Measurement of Input, Output, and Productivity,” *Econometrica*, 1982, 50, 1393–1414.

Clarke, Julian L. and Simon J. Evenett, “The Deterrent Effects of National Anticartel Laws: Evidence from the International Vitamins Cartel,” *Antitrust Bulletin*, 2003, 48, 689–727.

Conway, P. V. Janod and Giuseppe Nicoletti, “Product Market Regulation in OECD Countries: 1998 to 2003,” OECD Economics Department Working Paper, No. 568 2005.

– and – , “Product Market Regulation in the Non-Manufacturing Sectors of OECD Countries: Measurement and Highlights,” OECD Economics Department Working Paper, No. 530 2006.

Crandall, Robert W. and Clifford Winston, “Does Antitrust Policy Improve Consumer Welfare? Assessing the Evidence,” *Journal of Economic Perspectives*, 2003, 17, 3–26.

- Cusack, Thomas and Susanne Fuchs**, "Documentation Notes for Parties, Governments, and Legislatures Data Set," mimeo, Wissenschaftszentrum Berlin für Sozialforschung 2002.
- Denicolò, Vincenzo and Piercarlo Zanchettin**, "Competition, Market Selection and Growth," *The Economic Journal*, 2010, 120, 761–785.
- Disney, Richard, Jonathan Haskel, and Ylva Heden**, "Restructuring and Productivity Growth in UK Manufacturing," *Economic Journal*, 2003, 113, 666–694.
- Djankov, Simeon, Edward L. Glaeser, Rafael La Porta, Florencio Lopez de Silanes, and Andrei Shleifer**, "The New Comparative Economics," *Journal of Comparative Economics*, 2003, 31, 595–619.
- , **Rafael La Porta, Florencio Lopez de Silanes, and Andrei Shleifer**, "Courts," *Quarterly Journal of Economics*, 2003b, 118, 453–517.
- Duso, Tomaso and Jo Seldeslachts**, "The Political Economy of Mobile Telecommunications Liberalization: Evidence from the OECD Countries," *Journal of Comparative Economics*, 2010, 38, 199–216.
- **and Lars-Hendrik Röller**, "Endogenous Deregulation: Evidence from OECD Countries," *Economic Letters*, 2003, 81, 67–71.
- Dutz, Mark and Hayri Aydin**, "Does more intense competition lead to higher growth?," Discussion Paper 2249, CEPR 1999.
- **and Maria Vagliasindi**, "Competition Policy Implementation in Transition Economies: An Empirical Assessment," *European Economic Review*, 2000, 44, 762–772.
- Gilbert, Richard J.**, "Competition Policy for Intellectual Property," in Paolo Buccirossi, ed., *Handbook of Antitrust Economics*, MIT Press, 2008, pp. 519–542.
- Glaeser, Edward L., Rafael La Porta, Florencio Lopez de Silanes, and Andrei Shleifer**, "Do Institutions Cause Growth?," *Journal of Economic Growth*, 2004, 9, 271–303.

- Griffith, Rachel**, "Product Market Competition, Efficiency and Agency Costs: an Empirical Analysis," Working Papers, W01/12, IFS 2001.
- **and Rupert Harrison**, "The link Between Product Market Reform and Macroeconomic Performance," ECFIN economic papers No. 209, European Commission 2004.
- , – , **and Gareth Macartney**, "Product Market Reforms, Labour Market Institutions and Unemployment," *Economic Journal*, 2006, 117, C142–C166.
- , **Stephen Redding, and John van Reenen**, "Mapping the Two Faces of R&D: Productivity Growth in a Panel of OECD Industries," *Review of Economics and Statistics*, 2004, 86, 883895.
- Haskel, Jonathan**, "What Raises Productivity? The Microeconomics of UK Productivity Growth," Research Paper, Queen Mary, University of London 2000.
- Hausman, Jerry**, "Valuation of New Goods under Perfect and Imperfect Competition," in Timothy Bresnahan and Robert Gordon, eds., *The Economics of New Goods*, University of Chicago Press, 1997, pp. 209–437.
- Hicks, John R.**, "Annual Survey of Economic Theory: The Theory of Monopoly," *Econometrica*, 1935, 3, 1–20.
- Høj, Jens**, "Competition Law and Policy Indicators for the OECD countries," OECD Economics Department Working Papers No. 568 2007.
- Inklaar, Robert, Marcel P. Timmer, and Bart van Ark**, "Market Services Productivity," *Economic Policy*, 2008, 23, 141–194.
- Jorgenson, Dale W. and James A. Stephenson**, "Investment Behavior in United States Manufacturing, 1947-1960," *Econometrica*, 1967, 5, 169–220.
- , **Mun S. Ho, and Kevin J. Stiroh**, "Growth of U.S. Industries and Investments in Information Technology and Higher Education," in Carol Corrado, John Haltiwanger, and Daniel Sichel, eds., *Measuring Capital in the New Economy*, National Bureau of Economic Research Inc, 2005, pp. 403–478.

- Kee, Hiau Looi and Bernard Hoekman**, "Imports, Entry and Competition Law as Market Disciplines," *European Economic Review*, 2007, 51, 831–858.
- Klingemann, Hans-Dieter, Judith Bara, Ian Budge, Michael Macdonald, and Andrea Volkens**, *Mapping Policy Preferences II: Estimates for Parties, Electors and Governments in Central and Eastern Europe, European Union and OECD 1990-2003*, Oxford: Oxford University Press, 2006.
- Konings, Jozef, Patrick van Cayseele, and Frederic Warzynski**, "The Dynamics of Industrial Mark-ups in Two Small Open Economies: Does National Competition Policy Matter?," *International Journal of Industrial Organization*, 2001, 19, 841–859.
- Kovacic, William**, "Assessing the Quality of Competition Policy: The Case of Horizontal Merger Enforcement," *Competition Policy International*, 2009, 5, 129–150.
- La, Porta Rafael, Florencio Lopez de Silanes, and Andrei Shleifer**, "The Economic Consequences of Legal Origins," *Journal of Economic Literature*, 2008, 46, 285–332.
- , –, –, –, and **Robert W. Vishny**, "Legal Determinants of External Finance," *Journal of Finance*, 1997, 52, 1131–1150.
- , –, –, –, and –, "Law and Finance," *Journal of Political Economy*, 1998, 106, 1113–1155.
- , –, **Cristian Pop-Eleches, and Andrei Shleifer**, "Judicial Checks and Balances," *Journal of Political Economy*, 2004, 112, 445–470.
- Leibenstein, Harvey**, "Allocative Efficiency vs. "X-Efficiency"," *The American Economic Review*, 1966, 56, 392–415.
- Malmendier, Ulrike**, "Law and Finance at the Origin," *Journal of Economic Literature*, 2009, 47, 1076–1108.
- Nalebuff, Barry J. and Joseph E. Stiglitz**, "Information, Competition and Markets," *American Economic Review, Papers and Proceedings*, 1983, 73, 278–283.

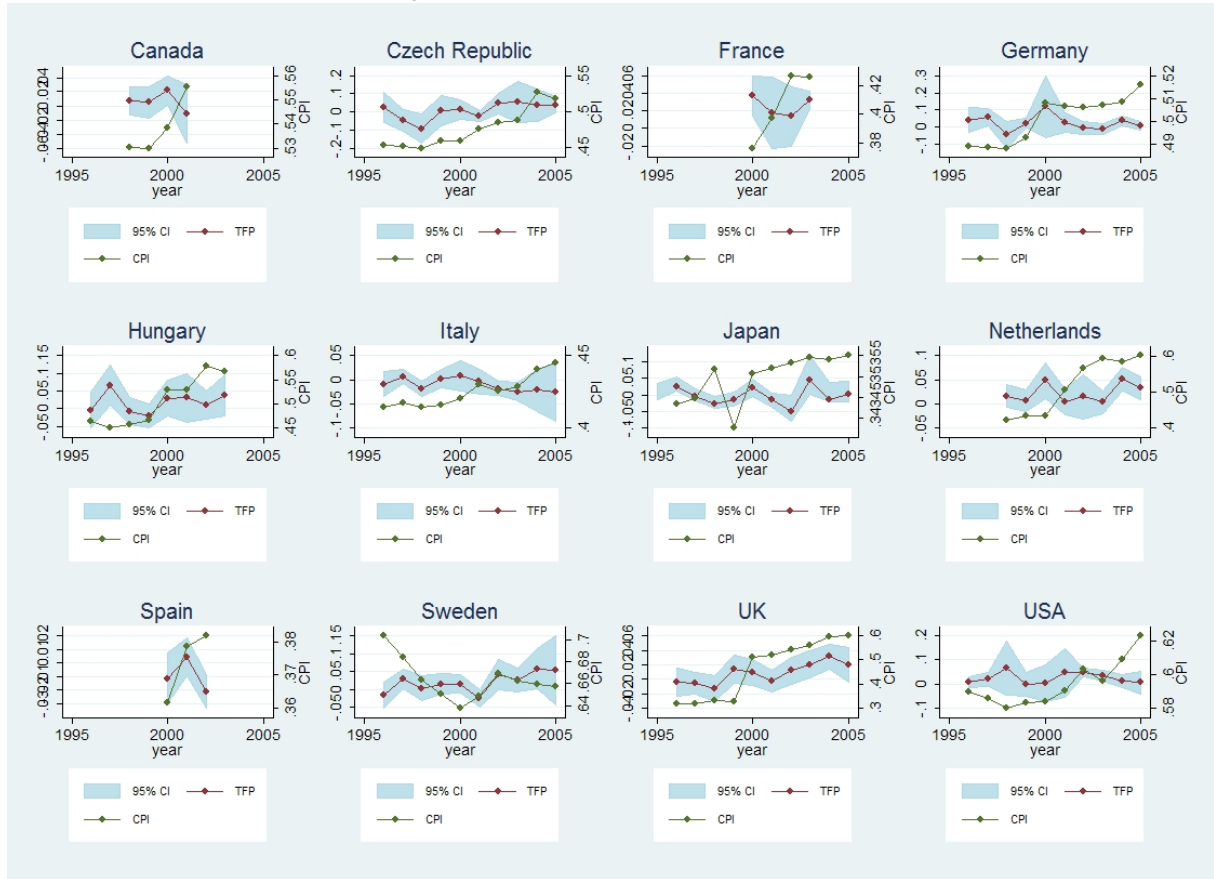
- Nickell, Stephen J.**, "Competition and Corporate Performance," *Journal of Political Economy*, 1996, 104, 724-746.
- , **Daphne Nicolitsas**, and **Neil Dryden**, "What Makes Firms Perform Well?," *European Economic Review*, 1997, 41, 783-796.
- Nicoletti, Giuseppe** and **Stefano Scarpetta**, "Regulation, Productivity and Growth: OECD Evidence," *Economic Policy*, 2003, 18, 9-72.
- Nocke, Volker**, "Collusion and Dynamic (Under-) Investment in Quality," *RAND Journal of Economics*, 2007, 38, 227-249.
- Paquet, Alain** and **Benoit Roubidoux**, "Issues on the Measurement of the Solow Residual and the Resting of its Exogeneity: Evidence for Canada," *Journal of Monetary Economics*, 2001, 47, 595-612.
- Sproul, Michael F.**, "Antitrust and Prices," *Journal of Political Economy*, 1993, 101, 741-754.
- Syverson, Chad**, "Product Substitutability and Productivity Dispersion," *Review of Economics and Statistics*, 2004, 86, 534-550.
- Timmer, Marcel P.**, **Mary O'Mahony**, and **Bart van Ark**, "EU KLEMS Growth and Productivity Accounts: An Overview," http://www.euklems.net/data/overview_07I.pdf 2007.
- Vickers, John**, "Concepts of Competition," *Oxford Economic Papers*, 1995, 47, 1-23.
- Voigt, Stefan**, "The Economic Effects of Competition Policy - Cross-Country Evidence Using Four New Indicators," *Journal of Development Studies*, 2009, 45, 1225-1248.
- von Mises, Ludwig**, "Human Action: A Treatise of Economics," <http://mises.org/Books/HumanActionScholars.pdf>, Ludwig von Mises Institute 1940.
- Werden, Gregory**, "The 1982 Merger Guidelines and the Ascent of the Hypothetical Monopolist Paradigm," *Antitrust Law Review*, 2003, 71, 253-269.
- Whinston, Michael D.**, *Lectures on Antitrust Economics*, Cambridge: MIT Press, 2006.

– **and Ilya R. Segal**, “Antitrust in Innovative Industries,” *American Economic Review*, 2007, 97, 1703–1730.

Woldendorp, Jaap, Hans Keman, and Ian Budge, *Party Government in 48 Democracies (1945-1998)*, Dordrecht: Kluwer Academic Publishers, 2000.

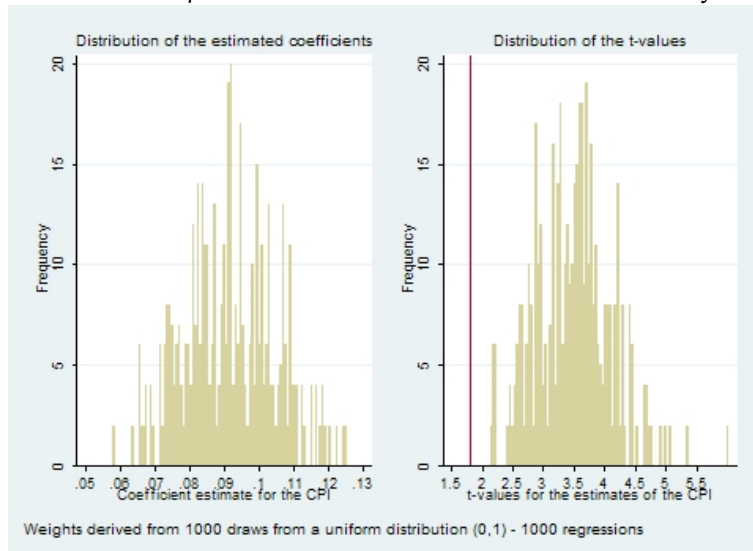
7 Figures and Tables

Figure 1: TFP Growth and the CPIs



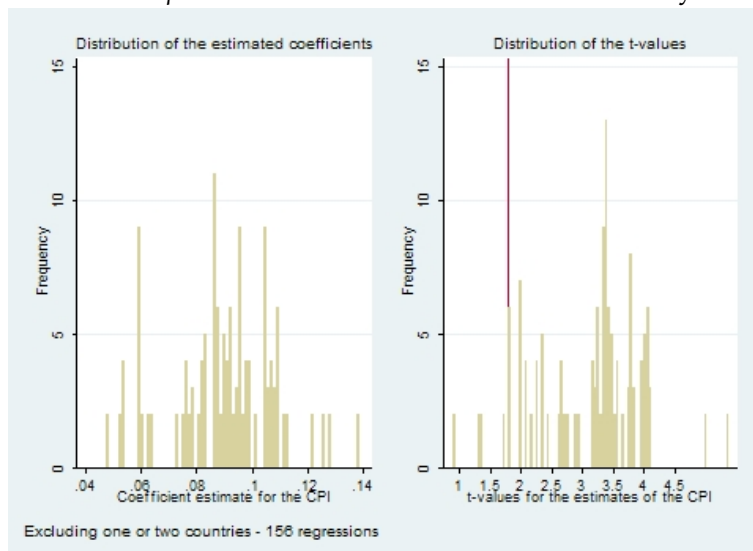
TFP growth is corrected for mark-ups. For each country, we report the weighted average of TFP growth across the 22 industries in the sample. The shaded area represents the 95% confidence interval around the mean.

Figure 2: Distribution of the β Coefficients and t -statistics obtained by Random Weights



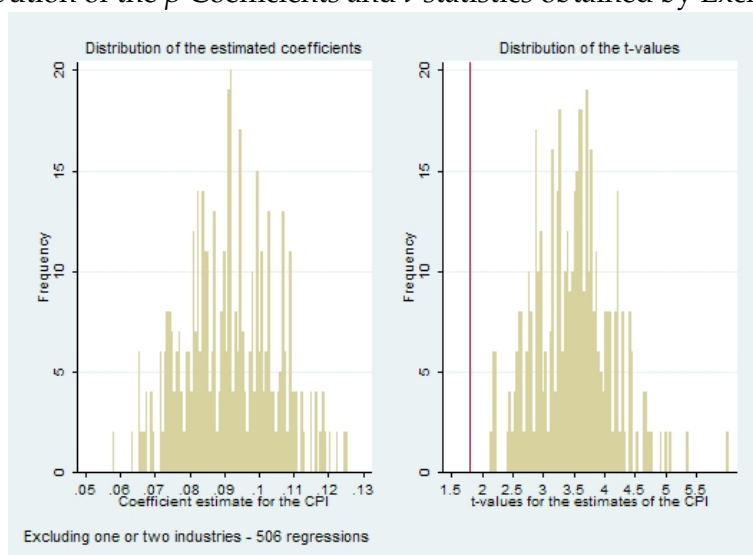
In the first panel, we represent the distribution of the estimated β coefficients from 1,000 regressions. In each of these regressions, the CPI index is built using random weights derived from a uniform distribution (0,1) and normalized to sum to 1. In the second panel, we represent the distribution of the t -statistics for the estimated coefficients. The red line represents the critical value for significance at the 10% level.

Figure 3: Distribution of the β Coefficients and t -statistics obtained by Excluding Countries



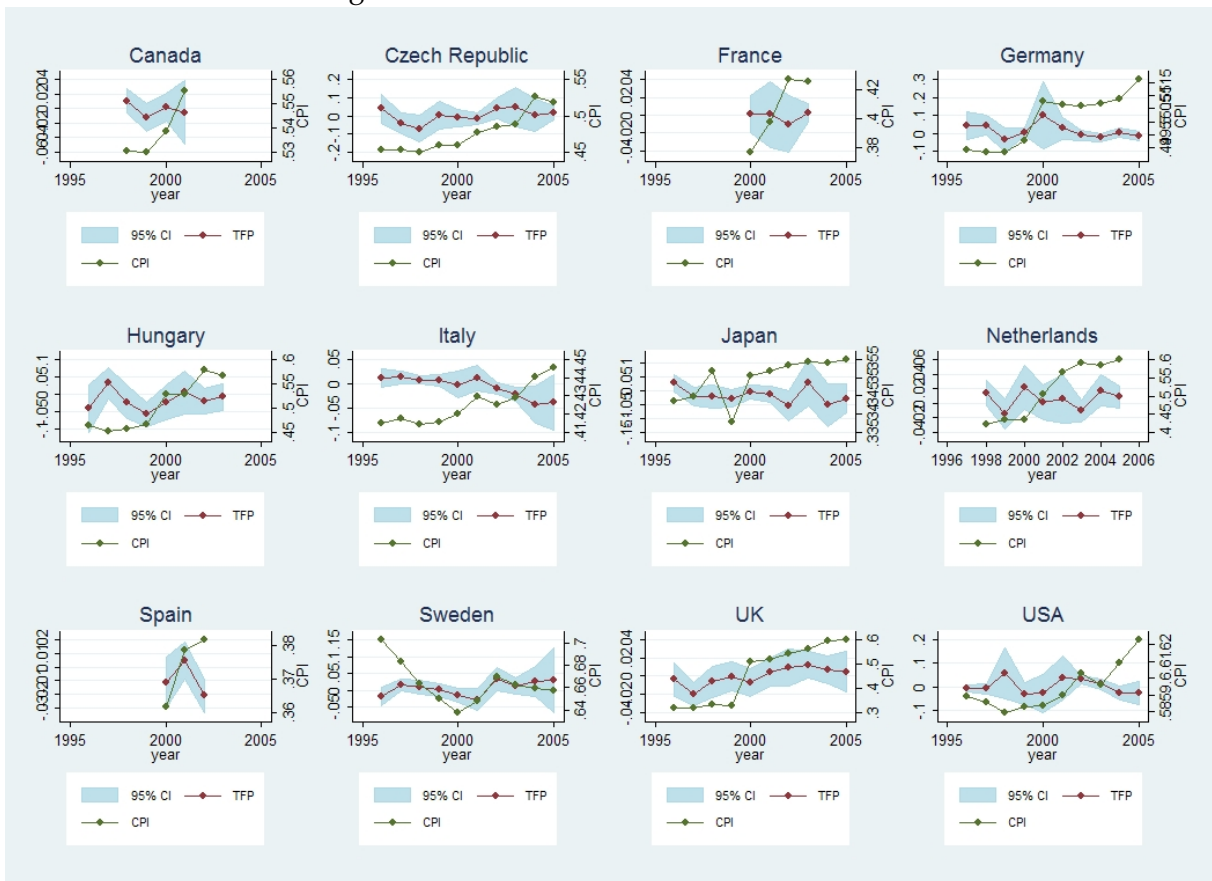
In the first panel, we represent the distribution of the estimated β coefficients from 156 regressions. In each of these regressions, we exclude one or two countries from our sample. In the second panel, we represent the distribution of the t -statistics for the estimated coefficients. The red line represents the critical value for significance at the 10% level.

Figure 4: Distribution of the β Coefficients and t -statistics obtained by Excluding Industries



In the first panel, we represent the distribution of the estimated β coefficients from 506 regressions. In each of these regressions, we exclude one or two industries from our sample. In the second panel, we represent the distribution of the t -statistics for the estimated coefficients. The red line represents the critical value for significance at the 10% level.

Figure 5: Residual TFP Growth and the CPIs



TFP growth is measured as the residual from equation (1), where we exclude the CPI from the regressors. The shaded area represents the 95% confidence interval around the mean TFP growth among the 22 industries for each country.

Table 1: Preliminary Statistics

| | Obs. | Mean | St. Dev. | Min. | Max |
|-----------------------------------|------|---------|----------|---------|---------|
| TFP Growth | 1847 | 0.0096 | 0.0686 | -0.2818 | 0.2727 |
| TFP Leader | 1847 | 0.0154 | 0.0931 | -0.7863 | 0.6246 |
| Technology Gap | 1847 | 0.6891 | 0.6697 | 0 | 5.6063 |
| R&D | 1463 | 0.0253 | 0.0574 | 0 | 0.4041 |
| Human Capital | 1783 | 0.1171 | 0.0977 | 0.0058 | 0.5588 |
| Trade openness | 1847 | 1.0096 | 1.8350 | 0 | 17.2785 |
| PMR | 1847 | 1.6721 | 0.5227 | 0.9234 | 3.0336 |
| CPI | 1847 | 0.4976 | 0.1019 | 0.3167 | 0.7035 |
| CPI.institution | 1847 | 0.6048 | 0.1114 | 0.3513 | 0.7735 |
| CPI.enforcement | 1847 | 0.2802 | 0.1587 | 0.0499 | 0.7513 |
| CPI.antitrust | 1847 | 0.5023 | 0.1032 | 0.3292 | 0.7047 |
| CPI.mergers | 1847 | 0.4834 | 0.1137 | 0.1372 | 0.6999 |
| Enforcement Costs | 1847 | 22.1471 | 8.2423 | 9.4000 | 33.5000 |
| Rule of Law | 1847 | 1.4263 | 0.4141 | 0.5251 | 1.8801 |
| Legal System | 1847 | 8.1494 | 1.0655 | 5.5667 | 9.6246 |
| Market regulation (per403) | 1847 | 1.3767 | 1.2564 | 0 | 5.5007 |
| Economic planning (per404) | 1847 | 0.3348 | 0.6229 | 0 | 2.6971 |
| Welfare state limitation (per505) | 1847 | 0.5264 | 0.5679 | 0 | 1.9637 |

We present preliminary statistics for all used variables in the selected estimation sample.

Table 2: Basic OLS Regressions - Aggregated Index

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) |
|----------------------|-----------------------|-----------------------|-----------------------|------------------------|-----------------------|-----------------------|------------------------|
| | OLS | OLS | OLS | OLS | OLS | OLS | OLS |
| TFP leader | 0.0653** (0.0233) | 0.0885*** (0.0251) | 0.0870*** (0.0257) | 0.0599** (0.0232) | 0.0811*** (0.0254) | 0.0863*** (0.0259) | 0.0651** (0.0228) |
| L.Techno Gap | 0.0075* (0.0041) | 0.0162** (0.00706) | 0.0168** (0.00724) | 0.0085* (0.0042) | 0.0181** (0.0069) | 0.0178** (0.0072) | -0.00169 (0.00566) |
| Industry trend | 0.0445*** (0.0052) | 0.127*** (0.0103) | 0.127*** (0.0100) | 0.0369*** (0.0052) | 0.131*** (0.0106) | 0.127*** (0.0100) | 0.0405*** (0.00595) |
| L.Import penetration | 0.0144*** (0.0040) | 0.0171** (0.0056) | 0.0174** (0.0056) | 0.0147*** (0.00415) | 0.0170** (0.0055) | 0.0171** (0.0056) | 0.0134*** (0.00400) |
| L.PMR | -0.0312 (0.0196) | -0.0380** (0.0172) | -0.0379** (0.0163) | -0.0390* (0.0205) | -0.0506** (0.0175) | -0.0410** (0.0168) | -0.0251 (0.0212) |
| L.CPI | 0.0924*** (0.0243) | 0.0827*** (0.0263) | 0.1064*** (0.0290) | 0.0945*** (0.0221) | 0.0800*** (0.0231) | 0.111*** (0.0291) | |
| L.CPI-low gap | | | | | | | 0.0548* (0.0304) |
| L.CPI-medium gap | | | | | | | 0.0821*** (0.0264) |
| L.CPI-high gap | | | | | | | 0.1223*** (0.0312) |
| L.R&D | | 0.6750*** (0.1880) | | | 0.6633** (0.2131) | | |
| L.Human Capital | | | | 0.286 (0.172) | 0.460* (0.218) | | |
| Constant | -0.137** (0.0536) | -0.433*** (0.0543) | -0.439*** (0.0516) | -0.00989 (0.0240) | 0.0147 (0.0292) | 0.0205 (0.0308) | -0.134** (0.0525) |
| R ² | 0.269 | 0.294 | 0.290 | 0.273 | 0.299 | 0.292 | 0.275 |
| Observations | 1847 | 1463 | 1463 | 1783 | 1408 | 1408 | 1847 |

The dependent variable is TFP growth corrected for mark-ups. Standard errors in parentheses are robust and allow for correlation among industries in the same country. In all regressions we insert country-industry dummies and time dummies. The symbols ***, **, and * represent significance at the 1%, 5%, and 10% significance respectively.

Table 3: OLS Regressions - Dissagregated Indexes

| | (1) OLS | (2) OLS | (3) OLS | (4) OLS |
|----------------------|-----------------------|-----------------------|-----------------------|-----------------------|
| TFP leader | 0.0656** (0.0233) | 0.0659** (0.0232) | 0.0654** (0.0233) | 0.0653** (0.0234) |
| Industry trend | 0.0428*** (0.0051) | 0.0438*** (0.0053) | 0.0444*** (0.0051) | 0.0443*** (0.0054) |
| L.Techno Gap | 0.0075* (0.0042) | 0.0076* (0.0042) | 0.0075* (0.0041) | 0.0075* (0.0042) |
| L.Import penetration | 0.0142*** (0.0040) | 0.0144*** (0.0040) | 0.0144*** (0.0040) | 0.0144*** (0.0040) |
| L.PMR | -0.0304 (0.0196) | -0.0266 (0.0250) | -0.0336 (0.0197) | -0.0249 (0.0206) |
| L.CPI.institution | 0.0705*** (0.0227) | | | |
| L.CPI.enforcement | | 0.0400* (0.0195) | | |
| L.CPI.antitrust | | | 0.0957*** (0.0255) | |
| L.CPI.mergers | | | | 0.0744*** (0.0221) |
| Constant | -0.133** (0.0551) | -0.117* (0.0594) | -0.132** (0.0526) | -0.143** (0.0587) |
| R^2 | 0.268 | 0.267 | 0.269 | 0.268 |
| Observations | 1847 | 1847 | 1847 | 1847 |

The dependent variable is TFP growth corrected for mark-ups. Standard errors in parentheses are robust and allow for correlation among industries in the same country. In all regressions we insert country-industry dummies and time dummies. The symbols ***, **, and * represent significance at the 1%, 5%, and 10% significance respectively.

Table 4: IV Regressions - Aggregated Index

| | (1) IV | (2) IV | (3) IV | (4) IV | (5) IV | (6) IV |
|--------------------------|------------------------|------------------------|------------------------|------------------------|-----------------------|------------------------|
| TFP leader | 0.0638*** (0.0186) | 0.0852*** (0.0211) | 0.0640*** (0.0186) | 0.0636*** (0.0186) | 0.0870*** (0.0210) | 0.0649*** (0.0185) |
| Industry trend | 0.0487** (0.0237) | 0.125*** (0.0398) | 0.0486** (0.0237) | 0.0491** (0.0238) | 0.126*** (0.0395) | 0.0459* (0.0236) |
| L.Techno Gap | 0.0074* (0.0040) | 0.0155*** (0.0055) | 0.0072* (0.0040) | 0.0074* (0.0040) | 0.0159*** (0.0055) | 0.0074* (0.0040) |
| L.Import penetration | 0.0146*** (0.0036) | 0.0175*** (0.0041) | 0.0146*** (0.00361) | 0.0147*** (0.00361) | 0.0173*** (0.0041) | 0.0145*** (0.0036) |
| L.R&D | | 0.481* (0.262) | | | 0.587** (0.260) | |
| L.PMR | -0.0402*** (0.0137) | -0.0543*** (0.0179) | -0.0493** (0.0195) | -0.0410*** (0.0142) | -0.0454** (0.0177) | -0.0388*** (0.0133) |
| L.CPI | 0.2220** (0.1020) | 0.2890** (0.1460) | 0.218** (0.102) | 0.233** (0.115) | 0.277** (0.143) | 0.136 (0.0832) |
| Constant | -0.276*** (0.0699) | -0.324*** (0.0780) | -0.258*** (0.0750) | -0.212** (0.105) | -0.0118 (0.0749) | 0.222*** (0.0799) |
| First-stage F-test (CPI) | 51.00 | 29.75 | 47.23 | 77.33 | 61.53 | 60.29 |
| First-stage F-test (PMR) | | | 194.49 | | | 147.84 |
| Sargan test | 2.616 (3) | 4.212 (3) | 2.450 (2) | 0.781 (1) | 0.899 (1) | 1.230 (2) |
| Wu-Hausman test | 0.2105 | 0.2219 | 0.4037 | 0.2366 | 0.5278 | 0.5067 |
| Observations | 1847 | 1463 | 1847 | 1847 | 1463 | 1847 |

The dependent variable is TFP growth corrected for mark-ups. Standard errors in parentheses are robust and allow for correlation among industries in the same country. The instruments in the IV regressions reported in columns 1, 2, and 3 are: coal, per108, per403, per404, per505. In column 1 only the CPI is instrumented, while in columns 2 and 3 both CPI and PMR are instrumented. The instruments in the IV regressions reported in columns 4, 5, and 6 are the average values of CPI and PMR among the other countries in the same group (European and non-European countries) and among the other countries in a different group. In columns 1, 2, 4, and 5 only the CPI is instrumented, while in columns 3 and 6 both CPI and PMR are instrumented. The value of the F-statistic for the test of excluded instruments in the first-stage regressions is reported. The Sargan statistic is distributed as a χ^2 and the degrees of freedom parameters are in parentheses. We report the p-value for the Wu-Hausman F-Statistic. In all regressions we insert country-industry dummies and time dummies. The symbols ***, **, and * represent significance at the 1%, 5%, and 10% significance respectively.

Table 5: First-Stage Regressions

| Specification | (1) | (2) | (3) | (4) | (5) | (6) |
|---|------------------------|------------------------|------------------------|------------------------|------------------------|------------------------|
| Dep. Var. | (1) | (3) | (3) | (4) | (6) | (6) |
| | CPI | CPI | PMR | CPI | CPI | PMR |
| L.per108 | 0.1292*** (0.0013) | 0.0124*** (0.0013) | -0.0071** (0.0030) | | | |
| L.per403 | -0.0083*** (0.0015) | -0.0126*** (0.0014) | 0.0578*** (0.0033) | | | |
| L.per404 | 0.0060** (0.0030) | 0.0034 (0.0054) | -0.0353*** (0.0072) | | | |
| L.per505 | 0.0011 (0.0039) | 0.0191*** (0.0031) | -0.2404*** (0.0075) | | | |
| L.CPI.NG | | | | -1.8651*** (0.1569) | -1.3312*** (0.1512) | 1.7896*** (0.2265) |
| L.CPI.G | | | | -0.2728*** (0.0335) | -0.2172*** (0.0328) | 0.0995** (0.0492) |
| L.PMR.NG | | | | | 0.0899 (0.0787) | -8.0118*** (0.1179) |
| L.PMR.G | | | | | -0.2287*** (0.0230) | -1.9799*** (0.0345) |
| Partial R^2 | 0.1148 | 0.1229 | 0.4056 | 0.0894 | 0.1641 | 0.7748 |
| Test of excluded instrum.: F(4,1574) | 51.00*** | 55.16*** | 286.56*** | 77.33*** | 77.23*** | 1354.06*** |
| Observations | 1847 | 1847 | 1847 | 1847 | 1847 | 1847 |

The dependent variable is CPI in columns 1, 2, 4, and 5 and PMR in columns 3 and 6. In column 1 and 4 only the CPI is instrumented, while in columns 2-3 and 5-6 both CPI and PMR are simultaneously instrumented. The Partial R-squared of excluded instruments and the value of the F-statistic for the test of excluded instruments in the first-stage regressions is reported. In all regressions we insert country-industry dummies and time dummies, as well as all the other exogenous variables from the main regression. The symbols ***, **, and * represent significance at the 1%, 5%, and 10% significance respectively.

Table 6: Interactions Regressions

| | (1) OLS | (2) OLS | (3) OLS | (4) OLS | (5) OLS | (6) OLS |
|---------------------|------------------------|-----------------------|-----------------------|-----------------------|-----------------------|-----------------------|
| Enforcement Cost | -0.0100*** (0.0007) | | -0.0063** (0.0027) | | | |
| Rule of law | 0.0211 (0.0298) | | | 0.0471 (0.0391) | | |
| Legal system | 0.0115* (0.0059) | | | | 0.0137* 0.0069 | |
| L.CPI | 0.0830*** (0.0204) | | | | | |
| L.CPI.LOe | | 0.0881*** (0.0143) | | | | |
| L.CPI.LOg | | 0.182*** (0.0324) | | | | |
| L.CPI.LOf | | 0.0206 (0.0406) | | | | |
| L.CPI.LOn | | 0.263** (0.117) | | | | |
| L.CPI.IEC | | | 0.240* (0.122) | | | |
| L.CPI.mEC | | | 0.110*** (0.0256) | | | |
| L.CPI.hEC | | | 0.0938** (0.0368) | | | |
| L.CPI.IRL | | | | 0.0837** (0.0310) | | |
| L.CPI.mRL | | | | 0.0945*** (0.0197) | | |
| L.CPI.hRL | | | | 0.117** (0.0532) | | |
| L.CPI.ILS | | | | | 0.0553 (0.0406) | |
| L.CPI.mLS | | | | | 0.0722*** (0.0253) | |
| L.CPI.hLS | | | | | 0.0830*** (0.0255) | |
| L.CPI.service | | | | | | 0.0091 (0.0501) |
| L.CPI.manufacturing | | | | | | 0.143*** (0.0420) |
| L.PMR.service | | | | | | -0.0485** (0.0189) |
| L.PMR.manufacturing | | | | | | -0.0235 (0.0188) |
| R ² | 0.273 | 0.270 | 0.271 | 0.270 | 0.270 | 0.272 |
| Observations | 1847 | 1847 | 1847 | 1847 | 1847 | 1847 |

The dependent variable is TFP growth corrected for mark-ups. Standard errors in parentheses are robust and allow for correlation among industries in the same country. In all regressions we insert country-industry dummies and time dummies. We control for the following variables 'TFP leader', 'Techno Gap', 'Industry trend', 'PMR', 'Import penetration' and a constant term but we do not report the coefficient estimates for space limitation and as they are comparable with those reported in Table 2. The symbols ***, **, and * represent significance at the 1%, 5%, and 10% significance respectively.

Table 7: Robustness Checks

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
|----------------------|-----------------------|-----------------------|-----------------------|-----------------------|-----------------------|-----------------------|----------------------|----------------------|-----------------------|
| | (OLS) | (OLS) | (OLS) | (OLS) | (OLS) | (OLS) | (OLS) | (OLS) | (OLS) |
| | Equal Weights | FA | EU | Non correct | LP | Long run I | Long run II | Long run III | Aggregated |
| TFP/LP leader | 0.0651** (0.0233) | 0.0657** (0.0232) | 0.0655** (0.0234) | 0.0372 (0.0340) | 0.0402 (0.0394) | 0.0734** (0.0248) | 0.0842 (0.272) | 0.0185 (0.139) | 0.2174** (0.0990) |
| L.Techno Gap | 0.0075* (0.0042) | 0.0075* (0.0041) | 0.0075* (0.0042) | 0.0564*** (0.0177) | 0.0084*** (0.0017) | -0.0027 (0.0056) | 0.0672** (0.0286) | -0.0070 (0.0152) | 0.0013 (0.0063) |
| Industry trend | 0.0464*** (0.0054) | 0.0426*** (0.0050) | 0.0450*** (0.0055) | 0.0533*** (0.0057) | 0.0536*** (0.0071) | 0.0548*** (0.0043) | 0.0078 (0.0265) | 0.0051** (0.0020) | 0.2531** (0.1025) |
| L.PMR | -0.0264 (0.0203) | -0.0315 (0.0200) | -0.0277 (0.0204) | -0.0141 (0.0213) | -0.0289 (0.0237) | 0.00642 (0.0353) | -0.171 (0.0969) | -0.0406 (0.0377) | -0.0125** (0.0058) |
| L.Import penetration | 0.0141*** (0.0039) | 0.0143*** (0.0040) | 0.0143*** (0.0039) | 0.0183*** (0.0052) | 0.0172*** (0.0040) | 0.00792 (0.0051) | 0.0812 (0.0506) | 0.0050* (0.0027) | 0.0044 (0.0041) |
| L.CPI | 0.0925*** (0.0209) | 0.0726** (0.0235) | 0.115*** (0.0369) | 0.0662* (0.0304) | 0.102*** (0.0298) | 0.0792* (0.0397) | 0.332* (0.156) | 0.0903* (0.0480) | 0.0417* (0.0236) |
| Constant | -0.161*** (0.0429) | -0.126** (0.0546) | -0.152** (0.0601) | -0.233*** (0.0430) | -0.644*** (0.0929) | -0.230*** (0.0628) | 0.0359 (0.182) | 0.0403 (0.0679) | -0.0024 (0.0135) |
| R ² | 0.269 | 0.268 | 0.268 | 0.274 | 0.302 | 0.301 | 0.414 | 0.394 | 0.272 |
| Observations | 1847 | 1847 | 1847 | 1850 | 1651 | 1275 | 1479 | 802 | 93 |

In all specifications we control for country-industry and time fixed-effects. In columns 1, 2, 3, 5, 6, 7, 8, and 9 the dependent variable is TFP growth corrected for mark-ups. In column 4 the dependent variable is TFP growth non-corrected for mark-ups. Column 1 and 2 report results for the model where the Aggregate CPI is constructed on the base of equal weights and the weights obtained by factor analysis (FA), respectively. Column 3 reports results for the model where the Aggregate CPI for EU member states incorporates information about EU competition policy. Column 4 reports results where all productivity measures are based on TFP non-corrected for mark-ups. Column 5 reports results where the technology gap and the productivity level of the country at the frontier are based on labor productivity. Column 6 reports results where all explanatory variables are lagged three years instead of one. Column 7 reports results based on a three-year time horizon; the explanatory variables are measured at the beginning of the period. Column 8 reports results based on a three-year time horizon; all variables are three-years averages. In this last specification, given the lack of degree of freedom, we use 12 country and 22 industry fixed effects, instead of 264 country-industry fixed-effects. Column 9 reports results based on country level observations; all industry variable are averaged using the industry value added as a weight. The symbols ***, **, and * represent significance at the 1%, 5%, and 10% significance respectively.

A The Indexes

The Competition Policy Indexes, CPIs, incorporate data on how the key features of a competition policy regime score against a benchmark of generally-agreed best practices and summarizes them. The CPIs have a pyramidal structure which encompasses a large number of sub-indicators that are progressively linearly combined using a set of weights at each level of aggregation. This structure is described in Tables A1, A2 and A3.

Table A1. The Low-level Indexes

| Abuses | Hard-core Cartels | Other agreements | Mergers |
|---|--|--|---|
| Independence: <i>Nature of prosecutor (1/2)</i> <i>Nature of adjudicator and role of government (1/2)</i> | Independence: <i>Nature of prosecutor (1/2)</i> <i>Nature of adjudicator and role of government (1/2)</i> | Independence: <i>Nature of prosecutor (1/2)</i> <i>Nature of adjudicator and role of government (1/2)</i> | Independence: <i>Nature of bodies involved in Phase 1 and 2 (1/2)</i> <i>Role of government in decision (1/2)</i> |
| Separation of powers: <i>Separation between adjudicator and prosecutor (2/3)</i> <i>Nature of appeal court (1/3)</i> | Separation of powers: <i>Separation between adjudicator and prosecutor (2/3)</i> <i>Nature of appeal court (1/3)</i> | Separation of powers: <i>Separation between adjudicator and prosecutor (2/3)</i> <i>Nature of appeal court (1/3)</i> | Separation of powers: <i>Separation between adjudicator and prosecutor (1/3)</i> <i>Separation between Phase 1 and 2 (1/3)</i> <i>Nature of appeal court (1/3)</i> |
| Quality of the law: <i>Standard of proof for predation and goals that inform decision (1/2)</i> <i>Standard of proof for refusal to deal and goals that inform decision (1/2)</i> | Quality of the law: <i>Standard of proof and goals that inform decision (1/2)</i> <i>Leniency program (1/2)</i> | Quality of the law: <i>Standard of proof for exclusive contracts and goals that inform decision</i> | Quality of the law: <i>Obligation to notify (1/2)</i> <i>Efficiency clause (1/2)</i> |
| Powers during investigation: <i>Combination of powers (3/4)</i> <i>Availability of interim measures (1/4)</i> | Powers during investigation: <i>Combination of powers</i> | Powers during investigation: <i>Combination of powers (3/4)</i> <i>Availability of interim measures (1/4)</i> | |
| Sanction policy and damages: <i>Sanctions to firms (1/3)</i> <i>Sanctions to individuals (1/3)</i> <i>Private actions (1/3)</i> | Sanction policy and damages: <i>Sanctions to firms (1/3)</i> <i>Sanctions to individuals (1/3)</i> <i>Private actions (1/3)</i> | Sanction policy and damages: <i>Sanctions to firms (1/3)</i> <i>Sanctions to individuals (1/3)</i> <i>Private actions (1/3)</i> | |
| Resources: <i>Budget (1/2)</i> <i>Staff (1/4)</i> <i>Staff skills (1/4)</i> | Resources: <i>Budget (1/2)</i> <i>Staff (1/4)</i> <i>Staff skills (1/4)</i> | Resources: <i>Budget (1/2)</i> <i>Staff (1/4)</i> <i>Staff skills (1/4)</i> | Resources: <i>Budget (1/2)</i> <i>Staff (1/4)</i> <i>Staff skills (1/4)</i> |
| | Sanctions and cases: <i>Number of cases opened (1/3)</i> <i>Max jail term imposed (2/3)</i> | | Cases: <i>Number of mergers examined</i> |

Table A1 shows the content of low-level indexes. The weights used to sum the information contained in each index are indicated in brackets.

Table A2 shows the eight medium-level indexes, which are given by the weighted average of the relevant low-level indexes. The weights are indicated in brackets.

Table A2. The medium-level Indexes

| | Abuses | Hard-core Cartels | Other agreements | Mergers |
|-------------------------------|---|---|---|--------------------------------------|
| Institutional features | <i>Independence</i> (1/6) | <i>Independence</i> (1/6) | <i>Independence</i> (1/6) | <i>Independence</i> (1/6) |
| | <i>Separation of powers</i> (1/6) | <i>Separation of powers</i> (1/6) | <i>Separation of powers</i> (1/6) | <i>Separation of powers</i> (1/3) |
| | <i>Quality of the law</i> (1/6) | <i>Quality of the law</i> (1/6) | <i>Quality of the law</i> (1/6) | <i>Quality of the law</i> (1/3) |
| | <i>Powers during investigation</i> (1/6) | <i>Powers during investigation</i> (1/6) | <i>Powers during investigation</i> (1/6) | |
| | <i>Sanctions and damages</i> (1/3) | <i>Sanctions and damages</i> (1/3) | <i>Sanctions and damages</i> (1/3) | |
| Enforcement features | <i>Resources</i> | <i>Resources</i> (2/3) | <i>Resources</i> | <i>Resources</i> (2/3) |
| | | <i>Cases</i> (1/3) | | <i>Cases</i> (1/3) |

Table A3 shows the different CPIs we built and the weights (in brackets) used in the aggregation process.

Table A3. The CPIs

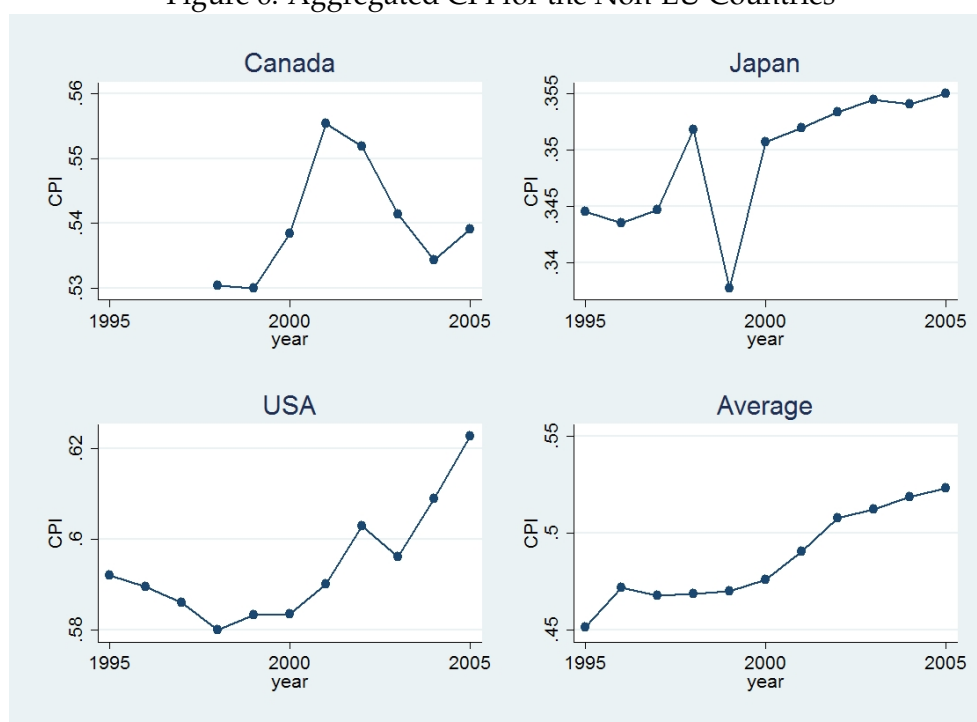
| The Aggregate CPI | | | | |
|--------------------------------|--|---|---|--|
| | The Antitrust CPI (3/4) | | | The Merger CPI (1/4) |
| | Hard-core Cartels (1/3) | Abuses (1/3) | Other agreements (1/3) | |
| Institutional CPI (2/3) | <i>Institutional features of hard core cartels</i> | <i>Institutional features of abuses</i> | <i>Institutional features of other agreements</i> | <i>Institutional features of hard core cartels</i> |
| Enforcement CPI (1/3) | <i>Enforcement features of hard core cartels</i> | <i>Enforcement features of abuses</i> | <i>Enforcement features of other agreements</i> | <i>Enforcement features of hard core cartels</i> |

We now turn to the values of the Aggregate CPIs for the countries in our sample over the period 1995-2005. Figures 6 to 8 give a general idea of the measure of the deterrence properties of the competition policy in those countries and of the relevant changes occurred over time. It is evident from them that there is substantial cross-sectional and cross-time variation. It should be stressed that the institutional component of the aggregate index takes a greater weight (2/3), hence the evolution of the Aggregate CPIs is mostly explained by the institutional features of

the competition policy which is relatively stable.⁵⁸

To allow a clearer interpretation of the results we include only a limited number of countries in each figure. Yet, to allow readers to easily perform comparisons among them, we report the sample average in each figure. Figure 6 shows the Institutional CPIs for the three OECD countries in our sample that are not part of the EU: Canada, Japan, and the US.

Figure 6: Aggregated CPI for the Non-EU Countries



As a starting point, the sample average of the aggregate CPIs shows an upwards trend during the sample period, which is common to almost all the 12 countries. Moreover, the time variation of the average index is significant with an average increase of almost 2% points per year (18% over the sample period). The Aggregate CPIs of the non-EU countries changed more or less markedly over the period under exam, and their levels differ considerably among each other. The aggregate CPI for the US takes very high values which are constantly among the highest in the sample ranging between 0.58 and 0.62, showing therefore a significant time

⁵⁸The enforcement features undergo more frequent changes and so do the Enforcement CPIs. For the sake of space we have only shown the values of Aggregate CPIs. For more details on the values of the other CPIs refer to Buccirosi et al. (2009a).

variation. The values for Canada are also quite high (between 0.53 and 0.56) and above the sample average. The range of variation is however limited to some percentage points per year. Japan's values are very low and among the lowest in the sample for the entire period (between 0.34 and 0.35). Differently from most other countries, also the changes in the Aggregate CPI are lower than an average of 1% per year. The reason behind Japan's low performance is manifold. First, Japan suffers from the lack of a leniency program for cartel whistleblowers. Second, in Japan there is no separation between the body that prosecutes violators of the antitrust law and the body that adjudicates such cases. Third, the Japanese CA has limited human and financial resources. Further elements are the absence of the possibility to start a class action and the fact that the Japanese competition legislation envisages the consideration of non strictly-economic goals when assessing the effects of abuses of dominance.

Figure 7 depicts the Aggregate CPIs for the large EU member states in our sample: France, Germany, Italy, Spain, and the UK.

Figure 7: Aggregated CPI for the Large EU Countries



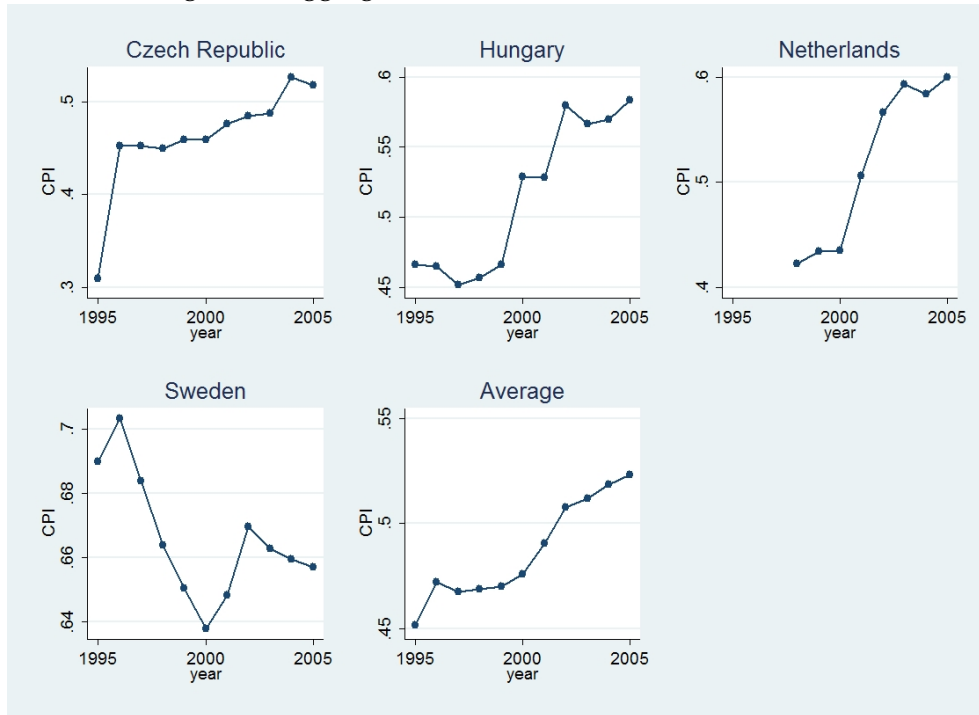
The first noticeable element in this figure is that the data for the first five years in the sample

are missing for Spain and France. This lack of information does not allow one to have a clear picture of the trend for these two jurisdictions. Anyhow, the Aggregate CPIs for these two countries, as well as for Italy, are very low and consistently below the sample average (0.38-0.42 for France, 0.36-0.42 for Spain, and 0.41-0.44 for Italy). Both Spain and France experience a substantial improvement between 2000 and 2003. The former benefited from the introduction of class action in 2001 and of the powers to investigate business premises in 2003. In the latter, the quality of the institutional CPI improved because of the introduction of a leniency program for cartels whistleblowers and of the obligation to notify mergers. Germany shows a good and constant performance ranging between 0.49 and 0.52. Notably, the CPIs for the UK start well below all the values of the CPIs of the other countries (0.3), but over time they become the highest in the group (0.6). This is due to the dramatic institutional changes that accompanied the introduction of the Competition Act in 2000, coupled with a steady increase in the financial and human resources of the two CAs.

Figure 8 depicts the Aggregate CPIs for the small EU member states in our sample: the Czech Republic, Hungary, the Netherlands, and Sweden.

Sweden is consistently the country with the highest CPI value, not just in this group but in the whole sample, yet this slowly declines over time (from 0.7 to 0.66) because of a reduction, in real terms, of the financial and human resources available to its CA. Instead, the CPIs for the other jurisdictions start below the sample average, but they all improve over time. The Czech Republic experiences a first, considerable shift in 1996, due to the CA acquiring independence from the government – previously all decisions were taken by a ministerial department. A further improvement takes place in 2004, when the power to investigate business premises is introduced. In the sample period, the CPI increases by 70% from a low of 0.3 to a high of 0.51. In Hungary the major changes happen in 2000, when there is an increase in the investigative powers of the CA and a shift in the criterion used to set the sanctions for antitrust infringements, which changed from a discretionary decision left to the adjudicator to an approach based on the firm's turnover. Moreover, in 2002 there was a substantial increase in the budget of the CA. These changes are captured by an increase in the CPI by over 30% from a low of 0.45 to a high of 0.59. The Netherlands did not have a CA before 1998. Hence, it was not possible to

Figure 8: Aggregated CPI for the Small EU Countries



calculate a CPI until that year. In subsequent years, the index steadily rises by almost 50% over the sample period from a low of 0.4 to a high of 0.6 as a consequence of a regular increase in the amount and in the quality of its CA's resources.

These three figures give a general idea of the factors that affect the ability of a competition policy regime to deter anti-competitive behavior in the jurisdictions included in our sample and of how these have changed over time. It is evident from them that there is substantial cross-sectional and cross-time variation.

Table A4 instead shows the ranking of the 12 countries in our sample based on the average value of their Aggregate CPIs over the years 1995 to 2005 and on its value in 2005. Sweden and the US are the best-scoring countries and this is true for each year in the sample, similarly France, Spain, and Japan constantly have the lowest scores. The UK and Canada are the countries that experience the most marked change.

Table A4 : The Ranking of the Countries on the Basis of the Aggregate CPIs

| Country | Ranking based on average score | Ranking based on 2005 score |
|----------------|--------------------------------|-----------------------------|
| Sweden | 1 | 1 |
| US | 2 | 2 |
| Canada | 3 | 6 |
| Netherlands | 4 | 3 |
| Hungary | 5 | 5 |
| Germany | 6 | 8 |
| Czech Republic | 7 | 7 |
| UK | 8 | 4 |
| Spain | 9 | 11 |
| Italy | 10 | 9 |
| France | 11 | 10 |
| Japan | 12 | 12 |

B The TFP Measures

In this appendix we describe in more detail the TFP growth and Technology Gap variables employed in our regressions.

TFP growth. The measure of TFP growth employed in our regressions is taken from the EU-KLEMS database.⁵⁹ The database improves substantially on the existing industry level databases, among which the OECD STAN database and its predecessor the ISDB database. The main limitation of previously existing databases is that they provide industry-level series on output, aggregate hours worked and aggregate capital stock, ignoring changes in the composition of factor inputs. As a result, TFP measures based on these aggregate quantities might be biased. On the contrary, the KLEMS database takes into account changes in the composition of the labor force over time. Furthermore, it discriminates among different types of capital input measures.

The TFP measure reported by the KLEMS database and employed in our regressions is based on the growth accounting methodology, which essentially consists of decomposing output growth into the contribution of input growth (labor and capital) and TFP growth.⁶⁰ TFP

⁵⁹The EU-KLEMS database is the result of a research project funded by the European Commission that involves major national level economic and statistical research centers. Details about the EU-KLEMS project can be found at the website: www.euklems.net. An overview of the methodology employed to collect data and build the measures of productivity can be found in Timmer et al. (2007).

⁶⁰The growth accounting methodology for computing productivity has a long standing history. For a full de-

measures within the growth accounting framework are based on several assumptions: in particular, it is assumed that markets are perfectly competitive and that inputs are fully utilized. Under these assumptions, TFP growth can be written as follows:

$$\Delta TFP_{ijt} = \ln\left(\frac{Y_{ijt}}{Y_{ijt-1}}\right) - \frac{1}{2}(\alpha_{ijt} + \alpha_{ijt-1})\ln\left(\frac{L_{ijt}}{L_{ijt-1}}\right) - \left(1 - \frac{1}{2}(\alpha_{ijt} + \alpha_{ijt-1})\right)\ln\left(\frac{K_{ijt}}{K_{ijt-1}}\right) \quad (2)$$

where Y_{ijt} is real value added, L_{ijt} measures the labor input and the K_{ijt} capital input. Within the EU-KLEMS database, accurate measures of labor and capital input are based on a breakdown of aggregate hours worked and aggregate capital stock into various components. Hours worked are cross-classified by various categories to account for differences in the productivity of various labor types, such as high- versus low-skilled labor. Similarly, capital stock measures are broken down into stocks of different asset types.⁶¹ The term α_{ijt} measures the labor share in value added. For our study, given that we measure the effectiveness of competition policy in promoting competition and ultimately efficiency, the main concern related to the TFP measure reported in the EU-KLEMS database is the assumption of perfect competition in the product markets. In order to take the existence of imperfectly competitive product markets into account, we modify the expression in equation (2) and multiply the labor share by industry-specific mark-ups.⁶²

We estimate industry level mark-ups as in Griffith and Harrison (2004), using the following equation:

$$Markup_{ijt} = \frac{ValueAdded_{ijt}}{LaborCosts_{ijt} + CapitalCosts_{ijt}} \quad (3)$$

where $ValueAdded_{ijt}$ is nominal value added, Labor Costs is labor compensation and Capital Costs is capital compensation.⁶³ The main source of data for computing mark-ups is still

description of the methodology see Jorgenson et al. (1967, 2005) and Caves (1982a).

⁶¹The EU-KLEMS database covers all the countries involved in our study except for Canada. For measuring TFP growth for Canada, we use data from the Groningen Growth and Development Centre (GGDC). The GGDC methodology is totally analogous to the one adopted by the EU-KLEMS consortium, of which the GGDC is member. The correlation between the EU-KLEMS TFP and the GGDC TFP is high (0.7) and strongly significant. However, we run specifications excluding Canada and results remain qualitatively and quantitatively unchanged.

⁶²In this, we follow the existing literature that explores the determinants of TFP growth. See, for example, Griffith et al. (2004), Aghion et al. (2009) and Nicoletti and Scarpetta (2003).

⁶³The Capital Costs measure is obtained by multiplying the capital stock for the user cost of capital, which takes into account the real interest rate and the extent of capital depreciation. For details see Griffith et al. (2006).

the EU-KLEMS database.⁶⁴ An important aspect to notice is that the measure of capital input necessary to compute capital costs is a somewhat cruder measure than the one employed in the construction of the TFP measure. In particular, we use an aggregate measure of capital stock, not accounting for different types of capital assets.⁶⁵ This capital stock measure is computed starting from the real gross fixed-capital formation series available in the EU-KLEMS database, using the perpetual inventory method.

Technology gap. One of the main regressors in our specifications is the technology gap between a country-industry in a given year and the technological frontier. There are several ways which can potentially be used to measure the technology gap. In our study, we follow the existing literature and use the TFP level to compute the distance to the technological frontier.⁶⁶ The computation of the technology gap is made in two steps. The first step consists of evaluating the level of TFP in each country-industry relative to a common reference point – the geometric mean of the TFPs of all other countries in the same industry. This measure of the TFP level with respect to the average is given by:

$$TFP_{ijt} = \ln\left(\frac{Y_{ijt}}{\bar{Y}_{jt}}\right) - \tilde{\sigma}_{ijt}\ln\left(\frac{L_{ijt}}{\bar{L}_{jt}}\right) - (1 - \tilde{\sigma}_{ijt})\ln\left(\frac{K_{ijt}}{\bar{K}_{jt}}\right)$$

where the output and input measures are the same employed in the measurement of TFP growth, and the bar denotes a geometric mean.⁶⁷ The variable $\tilde{\sigma}_{ijt} = \frac{1}{2}(\alpha_{ijt} + \bar{\alpha}_{jt})$ is the average of the labor share in country i and the geometric mean labor share. The technology leader is defined as the country-industry with the highest value for the TFP level relative to the common reference point. The second step for computing the technology gap consists of subtracting TFP_{ijt} from TFP_{Ljt} , where the latter is the TFP level in the identified country-industry leader. The technology gap variable used in our regressions is thus: $TechnoGap_{ijt} = TFP_{Ljt} - TFP_{ijt}$

⁶⁴For the computation of capital costs, we needed data on the inflation rate as well as on the yield on 10-years Federal Reserve Bonds. These come from the OECD MEI (Main Economic Indicators) database.

⁶⁵The reason why we use an aggregate measure of the capital stock is that the series on gross fixed-capital formation disaggregated for different types of assets are publicly available in the EU-KLEMS database only for a limited number of countries.

⁶⁶In the effort to verify the robustness of our results, we also employ a different measures of technology gap, based on labor productivity (value added per worker) differences among country-industries. The results remain basically unchanged, suggesting a stronger role for the technology gap in explaining TFP performance and weaker one for TFP growth on the technological frontier.

⁶⁷Data are aggregated using national level purchasing power parities (PPPs). For the base year we use for measuring real variables (2000), neither industry level PPPs for value added nor capital specific PPPs are available.

C The Assumptions on the Error Terms

Following the existing literature (e.g. Nicoletti and Scarpetta, 2003, Grififth et al., 2004, and Bourlès et al., 2010) we specified a particular structure for the individual effects and the error term in equation 1. In this appendix, we present and discuss a large amount of specifications, which are aimed at testing the robustness of our assumptions along two lines. First, since our data have a nested structure, as an industry is 'naturally' nested within a country, we follow Baltagi et al. (2001) and estimate several mixed-models to fit two-way, multilevel effects by maximum likelihood. Second, we more carefully analyze the autocorrelation structure of the residuals, to check and, eventually, correct for serial correlation in the residuals. Table A5 reports the results of our robustness checks.

We start by estimating a model with 12 country, 22 industry, and 9 time fixed-effect and cluster the standard error at the country level, which we use as a first benchmark (column 1). Then, we replicate our main specification with 264 country-industry and 9 time fixed-effects and standard error clustered at the country level (column 2). We then try a specification with country and time-industry fixed-effects (column 3). We then use three different specifications that make use of the nested structure we talked above and which are estimated by maximum likelihood with `xtmixed` in Stata. First, we specify country fixed-effects by the means of country dummies and use industry-within-country random effects. We allow for a complex, unspecified covariance structure and distinctly estimate all variances-covariances (column 4). We then assume country and industry-within-country random effects. Our model now has two random-effects equations. The first is a random intercept (constant-only) at the country level, the second a random intercept at the industry-within-country level (this, by the way, is exactly the model estimated by Baltagi et al. (2001) to investigate the productivity of public capital in private production). As before, we distinctly estimate all variances-covariances (column 5). While the size of the coefficient estimates is slightly affected, its sign and significance are not. In all specifications, we do find a strong and significant impact of the Aggregate CPI on TFP growth. Notice that, if we estimate a simple random effect model with country-industry random effects and time fixed-effects, i.e. a simplified version of specification (5), we also find a coefficient esti-

mate for the Aggregate CPI equal to 0.0550 and significant at the 1% level. However, when we run a Hausman test to verify whether the fixed or the random-effects specification should be preferred, we reject the appropriateness of the random-effects estimator.

The second robustness check concerns another aspect of the correlation structure of the residuals and, in particular, the potential existence of serial correlation. We start from our preferred fixed-effects specifications (1)-(3) with clustered standard errors at the country level. We run the Arellano and Bond (2001) test of autocorrelation of the first order.⁶⁸ The Arellano-Bond test rejects the null hypothesis in model (1) but not in model (2) and (3). We therefore re-estimate the basic models (1)-(3) by assuming a AR(1) structure for the error term. Results are reported in columns (6)-(8). Again, in all specifications we estimate a positive and significant coefficient for the CPI.⁶⁹ This is very similar in size to the coefficient estimated in our reference model. Eventually, the coefficients estimates are a bit larger in the models with AR(1) disturbances if compared to the basic specifications.

To conclude, while the structure for the error term that we adopted might appear to be subjective, we believe that it does not significantly affect our conclusions.

⁶⁸The test was originally proposed for a particular linear Generalized Method of Moments dynamic panel data estimator (Arellano and Bond, 1991), but is quite general in its applicability (more general than the xtserial test in Stata). It can be applied to linear GMM regressions in general, and thus to the special cases of ordinary least squares (OLS) and two-stage least-squares (2SLS). To run this test we therefore estimate the LSDV version of models (1)-(3).

⁶⁹Notice that the TFP level of the leader was dropped from specification (8) because of collinearity.

Table 8: Table A5 - Different Specifications with Various Individual Effects and Correlation Structures

| | (1) OLS | (2) OLS | (3) OLS | (4) MLE | (5) MLE | (6) OLS | (7) OLS | (8) OLS |
|---------------------------------|-------------------------|-----------------------|-------------------------|----------------------------|-----------------------------|------------------------|------------------------|-----------------------|
| TFP leader | 0.0738*** (0.0236) | 0.0653** (0.0233) | 0.133*** (0.0370) | 0.0869*** (0.0192) | 0.0990*** (0.0184) | 0.0728*** (0.0173) | 0.0657*** (0.0219) | dropped |
| L.Techno Gap | 0.0072 (0.0060) | 0.0075* (0.0041) | 0.0149 (0.0107) | 0.0050* (0.0030) | 0.0031 (0.0027) | 0.0073** (0.0030) | 0.0063 (0.0051) | 0.0054 (0.0040) |
| L.Import penetration | 0.0047** (0.0019) | 0.0144*** (0.0040) | 0.0048* (0.0022) | 0.0045*** (0.0011) | 0.0046*** (0.0010) | 0.0050*** (0.0009) | 0.0172*** (0.0050) | 0.0049*** (0.0011) |
| L.PMR | -0.0328* (0.0162) | -0.0312 (0.0196) | -0.0331* (0.0167) | -0.0334** (0.0132) | -0.0079* (0.0047) | -0.0323*** (0.0115) | -0.0529*** (0.0156) | -0.00419 (0.00352) |
| L.CPI | 0.0868*** (0.0243) | 0.0924*** (0.0243) | 0.0832** (0.0278) | 0.0875** (0.0380) | 0.0540** (0.0223) | 0.0849** (0.0410) | 0.115** (0.0476) | 0.0563*** (0.0132) |
| Constant | 0.0472 (0.0328) | -0.152*** (0.0440) | 0.0272 (0.0463) | -0.0045 (0.0256) | -0.0091 (0.0161) | -0.0021 (0.0329) | -0.0125 (0.0265) | -0.0077 (0.0058) |
| Correlation Structure | Clustered Fixed (12) | Clustered | Clustered Fixed (11) | Unstructured Fixed (12) | Unstructured Random (12) | AR(1) Fixed (12) | AR(1) | AR(1) Fixed (12) |
| Country Effects | Fixed (22) | | | | | Fixed (22) | | |
| Industry Effects | | | | | | | | |
| Industry-within-country Effects | Fixed (9) | Fixed (264) | Fixed (9) | Random (256) | Random (256) | Fixed (9) | Fixed (264) | Fixed (9) |
| Time Effects | | Fixed (9) | Fixed (9) | Fixed (9) | Fixed (9) | Fixed (9) | Fixed (9) | Fixed (9) |
| Industry-Time Effects | | | Fixed (242) | | | | | Fixed (242) |
| Arellano-Bond test | 0.258 | 0.054 | 0.077 | | | | | |
| LR test vs. linear regression | 1847 | 1847 | 1847 | 21.16*** | 1847 | 1591 | 1591 | 1628 |
| Observations | 0.121 | 0.269 | 0.255 | | | | | |
| R ² | | | | | 18.5*** | | | |

The dependent variable is TFP growth corrected for mark-ups. Columns 1-3 report results of the fixed-effects specifications (column 1: country, industry, time fixed-effects; column 2 country- industry and time fixed-effects; column 3 country, industry-time fixed-effects) where the error terms are robust and clustered at the country level. Column 4 reports result for a the mixed-model with country fixed-effects and industry-within-country random effects and column 5 report the results for the mixed model with country and industry-within-country random effects. In columns 4 and 5 the covariance structure is unspecified and all variances-covariances are distinctly estimated. Column 6-8 reproduce the fixed-effects specifications with clustered standard errors reported in 1-3, which are augmented to allow for an AR(1) structure in the error term. The symbols ***, **, and * represent significance at the 1%, 5%, and 10% significance respectively. The p-value of the Arellano-Bond (1991) test for AR(1) is reported.